PASSIVE SENSOR IMAGING USING CROSS CORRELATIONS OF
NOISY SIGNALS IN A SCATTERING MEDIUM

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Abstract. It is well known that the travel time or even the full Green’s function between two
passive sensors can be estimated from the cross correlation of recorded signal amplitudes generated
by ambient noise sources. It is also known that the direction of the energy flux from the noise
sources affects the estimation of the travel time. Using the stationary phase method we show here
that the travel time can be effectively estimated when the ray joining the two sensors continues into
the noise source region. We extend this analysis to passive sensor imaging of reflectors with different
ambient noise source configurations by suitably migrating the cross correlations. If in addition there
is multiple scattering in the medium then reflectors can be imaged with passive sensor networks or
arrays by migrating suitable fourth-order cross correlations. Fourth-order cross correlations can also
be used with auxiliary passive sensors in order to enhance travel time estimation in a scattering
medium.

Key words. Travel time estimation, passive sensor imaging, noise sources, random media.

AMS subject classifications. 35R30, 35R60, 86A15, 78A46

1. Introduction. The travel time between two sensors in an inhomogeneous
medium can be estimated using the coherent signals emitted by an impulse point
source and recorded by them. It can also be estimated using the ambient incoherent
noise by computing the cross correlation of noisy signals recorded by the sensors. The
ambient noise is generated by sources that are randomly distributed in space and are
statistically stationary in time. The cross correlation of signal amplitudes contains
information about the Green’s function of the wave equation from which the travel
time can be obtained. The background propagation velocity can then be estimated
from the travel times between sensors in a network covering the region of interest.

The first application of this technique was carried out in seismology where the
sensors were seismic stations recording velocity fields, the noise sources came from the
nonlinear interaction of the ocean swell with the coast that generates surface waves
[34], and the goal was to obtain an estimate of the background velocity map of a large
part of the Earth. The idea of exploiting the ambient noise and using the cross corre-
lation of noisy signals to retrieve information about travel times was first proposed in
helioseismology and seismology [18, 24, 31]. It has been applied to background veloc-
ity estimation from regional to local scales [21, 32, 20], volcano monitoring [29, 12, 11],
and petroleum prospecting [17]. In randomly layered media, correlation methods for
imaging are analyzed in [19]. When the support of the random noise sources extends
over all space and they are uncorrelated, that is, their correlation is a delta function,
it has been shown that the derivative of the cross correlation of the recorded signals
is the symmetrized Green’s function between the sensors [26]. This is also true with
spatially localized noise source distributions provided the waves propagate within an
ergodic cavity [15, 16, 2]. At the physical level this result can be obtained in both
open and closed environments provided that the recorded signals are equipartitioned
[22, 37, 25, 33, 23]. In an open environment this means that the recorded signals
are an uncorrelated and isotropic superposition of plane waves of all directions. In
a closed environment it means that the recorded signals are superpositions of normal modes with random amplitudes that are statistically uncorrelated and identically distributed.

In many applications, however, the noise source distribution is spatially limited and the recorded signals are not equipartitioned. As a result, the waves recorded by the sensors are dominated by the flux coming from the direction of the noise sources. The cross correlations of the recorded signals depend on the orientation of these sensors relative to the direction of the energy flux. This affects significantly the quality of the estimate for the Green’s function. It is good when the line between the sensors is along the direction of the energy flux and bad when it is perpendicular to it [34].

In this paper we use the stationary phase method to analyze the estimation of the travel time between two sensors from the cross correlation of recorded noisy signals. This asymptotic approach is valid when the decoherence time of the noise sources is small compared to the travel time between the two sensors. Using this formulation we analyze the dependence of the estimate of the travel time on the orientation of the ray between the sensors to the direction of the energy flux from the noise sources.

We extend the stationary phase analysis to passive sensor imaging of reflectors with different noise source configurations. In the presence of reflectors, the cross correlations between any two sensors have, in addition to the main peaks at the travel times between them, other peaks at lag times related to travel times from the sensors to the reflectors. We analyze in detail the relation between the secondary peaks in the cross correlations and travel times between sensors and reflectors for different spatial noise source distributions. With this information we show how to image the reflectors by migrating truncated correlations. If we have sensor data both with and without the reflectors then we can migrate differences of the cross correlations. The resolution of the images depends in an essential way on the support of the ambient noise source distribution relative to both the locations of the sensors and the reflectors. These results are obtained in a medium with a deterministic, smoothly varying background.

If there is multiple scattering in the medium, as in a randomly inhomogeneous medium, reflectors can still be imaged with passive sensor networks or arrays by migrating suitable fourth-order cross correlations. In addition to the primary peak at the travel time between the sensors and the secondary peaks at the travel times between sensors and reflectors, the cross correlation function has a long oscillatory tail, a coda, due to the multiple scattering. This coda is difficult to distinguish from the fluctuations in the correlation due to the noise sources. Following the coherent interferometric imaging ideas developed for broadband deterministic pulses in [8], we form space-time local cross correlations of the cross correlations and migrate them. The space-time window over which these local fourth-order cross correlations are computed is determined adaptively as described in [8].

Using the stationary phase method we also show that the travel time between two sensors can be estimated even in unfavorable situations provided that special fourth-order correlation functions with auxiliary sensors are used. By unfavorable we mean that the main component of the energy flux from the noise sources is roughly perpendicular to the ray connecting the two sensors. It is essential that there is scattering in the medium in order that the fourth-order correlations can be used effectively for enhancing travel time estimation between the two sensors. This result explains analytically how the iterated cross correlation technique proposed by Campillo and Stehly [13] works.
The paper is organized as follows. In Sections 2-3 we present the approach to travel time estimation and imaging using cross correlations of noisy signals. In Sections 4-5 we analyze in detail the cross correlation technique. We study the relation between the cross correlation and the Green’s function in Section 4. In Section 5 we use the stationary phase method to analyze the estimation of the travel time from the cross correlation function. In Section 6 we show how this technique can be used for imaging reflectors using only passive sensors. We show how iterated cross correlations can be used for imaging in a scattering medium in Section 7, and for travel time estimation in Section 8.

2. Random sources, random scattering and wave cross correlations.

The wave fields that are recorded by various sensors and are then cross correlated have properties that are inherited from the nature of the sources that generate them as well as from the environment in which they propagate. Analytically the simplest case is that of wave fields due to random sources in a homogeneous medium, which we discuss in the next subsection. We then discuss wave field cross correlations in strongly and weakly scattering media in Subsection 2.2.

2.1. Wave cross correlations in a homogeneous medium with random sources. Let \( u(t, x_1) \) and \( u(t, x_2) \) denote the time-dependent wave fields recorded by two sensors at \( x_1 \) and \( x_2 \). Their cross correlation function over the time interval \([0, T]\) with time lag \( \tau \) is given by

\[
C_T(\tau, x_1, x_2) = \frac{1}{T} \int_0^T u(t, x_1)u(t + \tau, x_2)dt.
\] (2.1)

In a homogeneous medium, if the source of the waves is a space-time stationary random field that is also delta correlated in space and time then it has been shown \([33, 26]\) that

\[
\frac{\partial}{\partial \tau} C_T(\tau, x_1, x_2) \simeq G(\tau, x_1, x_2) - G(-\tau, x_1, x_2),
\] (2.2)

where \( G \) is the Green’s function. This approximate equality holds for \( T \) sufficiently large and provided some limiting absorption is introduced to regularize the integral. It is briefly analyzed mathematically in Subsection 4.3. The main point here is that the time-symmetrized Green’s function can be obtained from the cross correlation if there is enough source diversity. In this case the wave field at any sensor is equipartitioned, in the sense that it is a superposition of uncorrelated plane waves of all directions. We can recover in particular the travel time \( \tau(x_1, x_2) \) from the singular support of the cross correlation.

The configuration (Figure 2.1) in which the spatial support of the noise sources extends over all space is rarely encountered in practice. Significant departures from this ideal situation occur when limited spatial diversity of the sources introduces directivity into the recorded fields, which affects the quality of the estimate of the Green’s function. If, in particular, the source distribution is spatially localized, then the flux of wave energy is not isotropic, and the cross correlation function is not symmetric (Figure 2.2). In some situations it may be impossible to distinguish the coherent part of the cross correlation function, which contains information about the travel time (Figure 2.3). A mathematical analysis is given in Section 5 using the stationary phase method.
2.2. Wave cross correlations in a scattering medium. In the case of a spatially localized distribution of noise sources, directional diversity of the recorded fields can be enhanced if there is sufficient scattering in the medium. An ergodic cavity with a homogeneous interior is a good example (Figure 2.4, right): Even with a source distribution that has very limited spatial support, the reverberations of the waves in the cavity generate interior fields with high directional diversity [15, 2]. Multiple scattering of waves by random inhomogeneities can also lead to wave field equipartition if the transport mean free path is short compared to the distance from the sources.
to the sensors [30, 21, 27]. The transport mean free path is the propagation distance over which wave energy transport in a scattering medium is effectively isotropic. In such a scattering medium (Figure 2.4, left), the inhomogeneities can be viewed as secondary sources in the vicinity of the sensors.

If in a random medium the transport mean free path is short compared to the distance between the sensors, then the cross correlation function still gives an estimate of the Green’s function, which is itself random because of the medium. However, its coherent part that has information about the travel time is essentially unobservable. The travel time can be estimated in a random medium when the noise sources are spatially limited provided that (i) the transport mean free path is short compared to the distance between the sources and the sensors, and (ii) it is long compared to the distance between the sensors.

3. Background velocity estimation and imaging. In this section we consider the two main reasons for estimating travel times from cross correlations. The first one, to which most of the recent work in geophysics is devoted, concerns background velocity estimation and is presented in Subsection 3.1. The second one concerns imaging of reflectors using passive sensor networks and arrays. This is presented in Subsection 3.2. In Subsection 3.3 we discuss how iterated cross correlation techniques can be used in scattering media both for imaging and background velocity estimation.
3.1. Background velocity estimation. From the travel times between sensors in a network we can estimate tomographically the background propagation velocity \[3\]. In seismic imaging, velocity estimation is usually done by other methods \[35, 5\]. It is clear that the density and topology of the sensor network, as well as the quality of the travel times estimates, influence the spatial resolution and the accuracy of the reconstructed velocity map. Tomographic travel time velocity analysis was applied successfully for surface-wave velocity estimation in Southern California \[28\], in Tibet \[38\], and in the Alps \[34\].

3.2. Array and distributed passive sensor imaging. The main purpose of this paper is to show that travel time estimation using cross correlations can also be used for passive sensor imaging. Consider an array of passive sensors located at \((x_j)_{j=1,...,N}\), which record signals emitted by random sources, and point reflectors located at \((z_{r,j})_{j=1,...,N_r}\). The recorded signals are denoted \(u(t, x_j), j = 1, ..., N\). In imaging we want to estimate the locations of the reflectors from this data set. If the impulse response matrix \(P(x_j, x_l, t)\) of the sensor network is known, even partially, then the usual migration techniques \[5, 14\] that back propagate the impulse responses numerically in a fictitious medium give estimates of the locations of the reflectors. Knowing the impulse response matrix requires, however, an active, broadband sensor network. This is because each sensor must emit an impulse that propagates in the medium and is then recorded by all the sensors. In Section 6 we show how the impulse response matrix of a sensor network with passive sensors can be estimated from the matrix of cross correlations

\[
C_T(\tau, x_j, x_l) = \frac{1}{T} \int_0^T u(t, x_j)u(t + \tau, x_l)dt, \quad j, l = 1, ..., N. \tag{3.1}
\]

In the presence of reflectors, the structure of the cross correlation between two sensors is more complex because it has additional peaks at arrival times between the sensors and the reflectors. These peaks are, however, relatively weak. In order that they can be effective in migration imaging we must either mask the main peaks at the travel times between the sensors, or subtract them if we have available sensor cross correlations in the absence of the reflectors.

Background velocity estimation is a preliminary step in implementing travel time migration, which is done numerically in a background medium with the estimated velocity. It is important to have some knowledge of the background velocity of the medium so that the numerical back propagation of the cross correlations will focus with good resolution near the correct reflector locations.

3.3. Interferometric travel time estimation and imaging in scattering media with iterated cross correlations. Another contribution presented here, in Section 7, is the use of iterated cross correlations to image reflectors in a scattering medium. By a scattering medium we mean a large collection of small random scatterers which we do not want to image, but which generate multiply scattered waves. As a result, the structure of the Green’s functions and, therefore, that of the cross correlations is more complex. In addition to the singular components corresponding to the arrival of the direct waves and of the waves interacting with the reflectors, which we want to image, the Green’s functions and the cross correlations have long incoherent tails, or coda, that come from the multiply scattered waves.

When migrating these cross correlations numerically using the estimated background velocity, the resulting images are distorted and have a lot of speckles. The
reason for this is that the coda due to the multiple scattering is also migrated. It is not possible to remove the coda from the cross correlations because the singular components due to the reflectors are buried in it. A natural way to reduce the adverse effects of multiple scattering is to use iterated cross correlations, after removing the principal singular components corresponding to the arrival times between the sensors. In order to be effective, the calculation of the iterated cross correlations must be done over adaptively estimated space-time windows as is done in coherent interferometric imaging (CINT) [8]. Adaptive CINT was introduced for active array and distributed sensor imaging in a scattering medium when the impulse response matrix is known. In Section 7 we describe how CINT can be used for passive array imaging in a scattering medium.

Multiple scattering can be advantageous in estimating the travel time between sensors because it enhances the directional diversity of the wave fields. The enhanced diversity is, however, in the incoherent coda of the cross correlations. In order to exploit it we can use iterated cross correlations with auxiliary sensors as suggested by Campillo and Stehly [13]. This iterated cross correlation technique is analyzed in Section 8 using the stationary phase method.

4. Extracting the Green’s function from the cross correlation.

4.1. The wave equation with noise sources. We consider the solution \( u \) of the wave equation in a \( d \)-dimensional inhomogeneous medium:

\[
\frac{1}{c^2(x)} \frac{\partial^2 u}{\partial t^2} - \Delta_x u = n^\ell(t, x).
\]  

(4.1)

The term \( n^\ell(t, x) \) models a random distribution of noise sources. It is a zero-mean stationary (in time) Gaussian process with autocorrelation function

\[
\langle n^\ell(t_1, y_1)n^\ell(t_2, y_2) \rangle = F^\ell(t_2 - t_1)\Gamma(y_1, y_2).
\]  

(4.2)

Here \( \langle \cdot \rangle \) stands for statistical average with respect to the distribution of the noise sources. We assume that the decoherence time of the noise sources is much smaller than typical travel times between sensors. If we denote with \( \epsilon \) the (small) ratio of these two time scales, we can then write the time correlation function \( F^\ell \) in the form

\[
F^\ell(t_2 - t_1) = F\left(\frac{t_2 - t_1}{\epsilon}\right),
\]  

(4.3)

where \( t_1 \) and \( t_2 \) are scaled relative to typical sensor travel times. The Fourier transform \( \tilde{F}^\ell \) of the time correlation function is a nonnegative, even real-valued function. It is proportional to the power spectral density of the sources:

\[
\tilde{F}^\ell(\omega) = \epsilon \hat{F}(\epsilon \omega),
\]  

(4.4)

where the Fourier transform is defined by

\[
\tilde{F}(\omega) = \int F(t)e^{i\omega t} dt.
\]  

(4.5)

The spatial distribution of the noise sources is characterized by the auto-covariance function \( \Gamma \). It is the kernel of a symmetric nonnegative definite operator. For simplicity, we will assume that the process \( n \) is delta-correlated in space:

\[
\Gamma(y_1, y_2) = \theta(y_1)\delta(y_1 - y_2),
\]  

(4.6)
where $\theta$ characterizes the spatial support of the sources. It is possible to consider a more general form for the spatial auto-covariance function, as in [2], but this complicates the calculations without changing the results quantitatively.

4.2. Statistical stability of the cross correlation function. The stationary solution of the wave equation has the integral representation

$$u(t, x) = \int \int_{-\infty}^{t} n^\varepsilon(s, y) G(t - s, x, y) ds dy$$

$$= \int \int n^\varepsilon(t - s, y) G(s, x, y) ds dy,$$  \hspace{1cm} (4.7)

where $G(t, x, y)$ is the time-dependent Green’s function. It is the fundamental solution of the wave equation

$$\frac{1}{c^2} \frac{\partial^2 G}{\partial t^2} - \Delta_x G = \delta(t) \delta(x - y),$$

starting from $G(0, x, y) = \frac{\partial}{\partial t} G(0, x, y) = 0$ (and continued on the negative time axis by $G(t, x, y) = 0 \forall t \leq 0$).

The empirical cross correlation of the signals recorded at $x_1$ and $x_2$ for an integration time $T$ is

$$C_T(\tau, x_1, x_2) = \frac{1}{T} \int_0^T u(t, x_1) u(t + \tau, x_2) dt.$$  \hspace{1cm} (4.9)

It is a statistically stable quantity, in the sense that for a large integration time $T$, $C_T$ is independent of the realization of the noise sources. This is stated in the following proposition proved in Appendix A.

**Proposition 4.1.** 1. The expectation of $C_T$ (with respect to the distribution of the sources) is independent of $T$:

$$\langle C_T(\tau, x_1, x_2) \rangle = C^{(1)}(\tau, x_1, x_2),$$  \hspace{1cm} (4.10)

where $C^{(1)}$ is given by

$$C^{(1)}(\tau, x_1, x_2) = \int dy \int ds ds' G(s, x_1, y) G(\tau + s + s', x_2, y) F^\varepsilon(s') \theta(y),$$  \hspace{1cm} (4.11)

or equivalently by

$$C^{(1)}(\tau, x_1, x_2) = \int dy \int d\omega \hat{G}(\omega, x_1, y) \hat{G}(\omega, x_2, y) \hat{F}^\varepsilon(\omega) e^{-i\omega \tau} \theta(y).$$  \hspace{1cm} (4.12)

2. The empirical cross correlation $C_T$ is a self-averaging quantity:

$$C_T(\tau, x_1, x_2) \xrightarrow{T \to \infty} C^{(1)}(\tau, x_1, x_2),$$  \hspace{1cm} (4.13)

in probability with respect to the distribution of the sources. More precisely, the fluctuations of $C_T$ around its mean value $C^{(1)}$ are of order $T^{-1/2}$ for $T$ large compared to the decoherence time of the sources. The asymptotic covariance function of $C_T$ is given by (A.4).

Both expressions (4.11) and (4.12) will be used in the following sections.
4.3. Emergence of the Green’s function for an extended distribution of sources in a homogeneous medium. We give an elementary proof of the relation between the cross correlation and the Green’s function when the medium is homogeneous with background velocity $c_0$ and the source distribution extends over all space, i.e. $\theta \equiv 1$, as in Figure 2.1. In this case the signal amplitude diverges because the contributions from noise sources far away from the sensors are not damped. For a well-posed formulation we need to introduce some dissipation, so we consider the solution $u$ of the damped wave equation:

$$\frac{1}{c_0^2} \left( \frac{1}{T_a} + \frac{\partial}{\partial t} \right)^2 u - \Delta u = n^\varepsilon(t, x). \tag{4.14}$$

A somewhat different form of the following proposition, with delta-correlated in time sources and with a different definition of dissipation, can be found in [26].

**Proposition 4.2.** In a three-dimensional open medium with dissipation and if the source distribution extends over all space $\theta \equiv 1$, then

$$\frac{\partial}{\partial \tau} C^{(1)}(\tau, x_1, x_2) = -\frac{e_0^2}{4} \int \exp\left( -\frac{|x_1-x_2|^2}{e_0^2} \right) \left[ F^\varepsilon * G(\tau, x_1, x_2) - F^\varepsilon * G(-\tau, x_1, x_2) \right], \tag{4.15}$$

where $*$ stands for the convolution in $\tau$ and $G$ is the Green’s function of the homogeneous medium without dissipation:

$$G(t, x_1, x_2) = \frac{1}{4\pi|x_1-x_2|} \delta \left( t - \frac{|x_1-x_2|}{c_0} \right).$$

If the decoherence time of the sources is much shorter than the travel time (i.e., $\varepsilon \ll 1$), then $F^\varepsilon$ behaves like a Dirac distribution in (4.15) and we have

$$\frac{\partial}{\partial \tau} C^{(1)}(\tau, x_1, x_2) \simeq e^{-\frac{|x_1-x_2|^2}{e_0^2}} \left[ G(\tau, x_1, x_2) - G(-\tau, x_1, x_2) \right],$$

up to a multiplicative constant. It is therefore possible to estimate the travel time $\tau(x_1, x_2) = |x_1-x_2|/c_0$ between $x_1$ and $x_2$ from the cross correlation, with an accuracy of the order of the decoherence time of the noise sources.

**Proof.** The Green’s function of the homogeneous medium with dissipation is:

$$G_a(t, x_1, x_2) = G(t, x_1, x_2) e^{-\frac{\tau^2}{c_0}}.$$

The cross correlation function is given by (4.11):

$$C^{(1)}(\tau, x_1, x_2) = \int dy \int dsds' G_a(s, x_1, y) G_a(\tau + s + s', x_2, y) F^\varepsilon(s').$$

Integrating in $s$ and $s'$ gives

$$C^{(1)}(\tau, x_1, x_2) = \int \frac{dy}{16\pi^2 |x_1-y| |x_2-y|} e^{-\frac{|x_1-y|+|x_2-y|}{e_0^2}} F^\varepsilon \left( \tau - \frac{|x_1-y|-|x_2-y|}{c_0} \right).$$

We parameterize the locations of the sensors by $x_1 = (h, 0, 0)$ and $x_2 = (-h, 0, 0)$, where $h > 0$, and we use the change of variables for $y = (x, y, z)$:

$$\begin{cases} x = h \sin \theta \cosh \phi, & \phi \in (0, \infty), \\ y = h \cos \theta \sinh \phi \cos \psi, & \theta \in (-\pi/2, \pi/2), \\ z = h \cos \theta \sinh \phi \sin \psi, & \psi \in (0, 2\pi), \end{cases}$$

where $\theta$ and $\phi$ are polar angles and $\psi$ is an azimuthal angle. This transformation rotates the problem into a plane, where the Green’s function $G_a$ is expressed in terms of the polar coordinates $r^2 = x^2 + y^2 + z^2$ and $\theta$ and $\phi$.

$$G_a(r^2, \theta, \phi) = \frac{1}{r} \exp\left( -\frac{r^2}{c_0} \right).$$
whose Jacobian is \( J = h^3 \cos \theta \sinh \phi (\cosh^2 \psi - \sin^2 \theta) \). Using the fact that \( |x_1 - y| = h (\cosh \phi - \sin \theta) \) and \( |x_2 - y| = h (\cosh \phi + \sin \theta) \), we get

\[
C^{(1)}(\tau, x_1, x_2) = \frac{h}{8\pi} \int_0^\infty d\phi \sinh \phi \int_{-\pi/2}^{\pi/2} d\theta \cos \theta e^{-\frac{2h \cosh \phi}{c_0^2}} F^\varepsilon (\tau + \frac{2h \sin \theta}{c_0}).
\]

After the new change of variables \( u = h \cosh \phi \) and \( s = (2h/c_0) \sin \theta \), we obtain

\[
C^{(1)}(\tau, x_1, x_2) = \frac{c_0^2 T_\alpha}{32\pi h} e^{-\frac{2h}{c_0}} \int_{-2h/c_0}^{2h/c_0} F^\varepsilon (\tau + s) ds.
\]

By differentiating in \( \tau \), we get

\[
\frac{\partial}{\partial \tau} C^{(1)}(\tau, x_1, x_2) = \frac{c_0^2 T_\alpha}{32\pi h} e^{-\frac{2h}{c_0}} \left[ F^\varepsilon (\tau + \frac{2h}{c_0}) - F^\varepsilon (\tau - \frac{2h}{c_0}) \right],
\]

which is the desired result since \( |x_1 - x_2| = 2h \).

\[\square\]

### 4.4. Emergence of the Green's function for an extended distribution of sources in an inhomogeneous medium

The cross correlation function is closely related to the symmetrized Green's function from \( x_1 \) to \( x_2 \) not only for a homogeneous medium but also for inhomogeneous media, as discussed in the introduction. Here we give a simple and rigorous proof for an open inhomogeneous medium in the case in which the noise sources are located on the surface of a sphere that encloses both the inhomogeneous region and the sensors, located at \( x_1 \) and \( x_2 \). The proof is based on an approximate identity that follows from Green's identity and the Sommerfeld radiation condition. This approximate identity can be viewed as a version of the Helmholtz-Kirchhoff integral theorem (known in acoustics [4, p. 473] and in optics [10, p. 419]) and it is also presented in [36].

**Proposition 4.3.** Let us assume that the medium is homogeneous with background velocity \( c_e \) outside the ball \( B(0, D) \) with center \( 0 \) and radius \( D \). Then, for any \( x_1, x_2 \in B(0, D) \) we have for \( L \gg D \):

\[
\bar{G}(\omega, x_1, x_2) - \bar{G}(\omega, x_1, x_2) = \frac{2i\omega}{c_e} \int_{\partial B(0, L)} \bar{G}(\omega, x_1, y) \bar{G}(\omega, x_2, y) d\sigma(y).
\]

**Proof.** We consider the equation satisfied by the time-harmonic Green's function with the source at \( x_2 \) and the complex conjugate form of this equation with the source at \( x_1 \):

\[
\Delta y \bar{G}(\omega, y, x_2) + \frac{\omega^2}{c^2(y)} \bar{G}(\omega, y, x_2) = -\delta(y - x_2),
\]

\[
\Delta y \bar{G}(\omega, y, x_1) + \frac{\omega^2}{c^2(y)} \bar{G}(\omega, y, x_1) = -\delta(y - x_1).
\]

We multiply the first equation by \( \bar{G}(\omega, y, x_1) \) and subtract the second equation multiplied by \( \bar{G}(\omega, y, x_2) \):

\[
\nabla y \cdot \left[ \bar{G}(\omega, y, x_1) \nabla y \bar{G}(\omega, y, x_2) - \bar{G}(\omega, y, x_2) \nabla y \bar{G}(\omega, y, x_1) \right]
\]

\[
= \bar{G}(\omega, y, x_2) \delta(y - x_1) - \bar{G}(\omega, y, x_1) \delta(y - x_2)
\]

\[
= \bar{G}(\omega, x_1, x_2) \delta(y - x_1) - \bar{G}(\omega, x_1, x_2) \delta(y - x_2),
\]
where we have used the reciprocity property $\hat{G}(\omega, x_2, x_1) = \hat{G}(\omega, x_1, x_2)$. We next integrate over the ball $B(0, L)$ and use the divergence theorem:

$$\int_{\partial B(0, L)} \mathbf{n} \cdot \left[ \overline{\nabla \hat{G}(\omega, y, x_1)} \nabla \hat{G}(\omega, y, x_2) - \hat{G}(\omega, y, x_1) \nabla \overline{\hat{G}(\omega, y, x_2)} \right] d\sigma(y)$$

$$= \hat{G}(\omega, x_1, x_2) - \overline{\hat{G}(\omega, x_1, x_2)},$$

where $\mathbf{n}$ is the unit outward normal to the ball $B(0, L)$, which is $\mathbf{n} = y/|y|$. The Green’s function also satisfies the Sommerfeld radiation condition

$$\lim_{|y| \to \infty} \frac{|y|^{1/2}}{i} \left( \frac{y}{|y|} \cdot \nabla_y + i \frac{\omega}{c_0} \right) \hat{G}(\omega, y, x_1) = 0,$$

uniformly in all directions $y/|y|$. Under the conditions stated in the proposition, we can substitute $i(\omega/c_0)\hat{G}(\omega, y, x_2)$ for $\mathbf{n} \cdot \nabla \hat{G}(\omega, y, x_2)$ in the surface integral over $\partial B(0, L)$, and $-i(\omega/c_0)\overline{\hat{G}(\omega, y, x_1)}$ for $\mathbf{n} \nabla \overline{\hat{G}(\omega, y, x_1)}$, which gives the desired result.

Note that it is important that the medium be homogeneous in the exterior of the ball $B(0, D)$ in order to be able to use the radiation condition. \qed

The right side of the Helmholtz-Kirchhoff identity (4.16) is related to the representation (4.12) of the cross correlation function $C^{(1)}$ in the Fourier domain. Therefore, by substituting (4.16) into (4.12) we get the following corollary.

**Corollary 4.4.** We assume that

1) the medium is homogeneous with background velocity $c_0$ outside the ball $B(0, D)$ with center $0$ and radius $D$,

2) the sources are localized with a uniform density on the sphere $\partial B(0, L)$ with center $0$ and radius $L$.

If $L \gg D$, then for any $x_1, x_2 \in B(0, D)$

$$\frac{\partial}{\partial \tau} C^{(1)}(\tau, x_1, x_2) = -F^\varepsilon \ast G(\tau, x_1, x_2) + F^\varepsilon \ast G(-\tau, x_1, x_2), \quad (4.17)$$

up to a multiplicative factor. Here $\ast$ stands for convolution in $\tau$.

If in addition we have $\varepsilon \ll 1$, then $F^\varepsilon$ behaves approximately like a delta distribution acting on the Green’s function and we get the form (2.2).

In proving the identity (2.2) in the previous subsections it was necessary that the noise source distribution generate an isotropic field at the sensors, coming from all directions. In the next sections we will consider cases in which this isotropy condition is not fulfilled.

5. **Stationary phase analysis for travel time estimation with spatially localized noise sources.** We study in this section the cross correlation function when the support of the sources is spatially limited. We assume that the fluctuations of the medium parameters are modeled by a smooth background velocity profile $c_0(x)$ which is homogeneous outside a large sphere that encloses the sensors and the sources. The outgoing time-harmonic Green’s function $G_0$ of the medium is the solution of

$$\Delta_x \hat{G}_0(\omega, x, y) + \frac{\omega^2}{c_0^2(x)} \hat{G}_0(\omega, x, y) = -\delta(x - y), \quad (5.1)$$

along with the radiation condition at infinity. When the background is homogeneous with constant wave speed $c_0$ and wavenumber $k = \omega/c_0$, then the homogeneous out-
going time-harmonic Green’s function is

\[ \hat{G}_0(\omega, x, y) = \frac{e^{ik|x-y|}}{4\pi|x-y|} \quad (5.2) \]

in three-dimensional space, and

\[ \hat{G}_0(\omega, x, y) = i\frac{4}{4} H_0^{(1)}(k|x-y|) \quad (5.3) \]

in two-dimensional space. Here \( H_0^{(1)} \) is the zeroth order Hankel function of the first kind. Using the asymptotic form of the Hankel function \[1, \text{ formula 9.2.3}\], we see that the high-frequency behavior of the Green’s function is related to the homogeneous medium travel time \(|x - y|/c_0\):

\[ \hat{G}_0(\omega \varepsilon, x, y) \sim \frac{1}{|x-y|^{(d-1)/2}} e^{i\frac{\pi}{2}(x-y)} \quad (5.4) \]

which is valid when \( \varepsilon \ll 1 \). Here the coefficients \( a(x, y) \) and \( \tau(x, y) \) are smooth except at \( x = y \). The amplitude \( a(x, y) \) satisfies a transport equation and the travel time \( \tau(x, y) \) satisfies the eikonal equation. It is a symmetric function \( \tau(x, y) = \tau(y, x) \) and it can be obtained from Fermat’s principle

\[ \tau(x, y) = \inf \left\{ T \, \text{ s.t. } \exists (X_t)_{t \in [0, T]} \in C^1, X_0 = x, X_T = y, \left| \frac{dX_t}{dt} \right| = c_0(X_t) \right\} \quad (5.5) \]

A curve \((X_t)_{t \in [0, T]}\) that produces the minimum in (5.5) is a ray and it satisfies Hamilton’s equations (see Appendix B). For simplicity we assume that the background speed \( c_0(x) \) is such that there is a unique ray joining any pair of points \((x, y)\) in the region of interest.

We can now describe the behavior of the cross correlation function between \(x_1\) and \(x_2\) when \( \varepsilon \) is small, with and without directional energy flux from the sources.

**Proposition 5.1.** As \( \varepsilon \) tend to zero, the cross correlation \( C^{(1)}(\tau, x_1, x_2) \) has singular components if and only if the ray going through \( x_1 \) and \( x_2 \) reaches into the source region, that is, into the support of the function \( \theta \). In this case there are either one or two singular components at \( \tau = \pm \tau(x_1, x_2) \).

More precisely, any ray going from the source region to \( x_2 \) and then to \( x_1 \) gives rise to a singular component at \( \tau = -\tau(x_1, x_2) \). Rays going from the source region to \( x_1 \) and then to \( x_2 \) give rise to a singular component at \( \tau = \tau(x_1, x_2) \).

This proposition explains why travel time estimation is bad when the ray joining \( x_1 \) and \( x_2 \) is roughly perpendicular to the direction of the energy flux from the noise sources, as in the middle of Figure 5.1.

The stationary phase contributions to the singular components of the cross correlation come from pairs of ray segments. The first ray goes from a source point to \( x_2 \) and the second ray goes from the same source point and with the same initial angle
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Fig. 5.1. If the ray going through \( x_1 \) and \( x_2 \) (solid line) enters into the source region (left figure), then the travel time can be estimated from the cross correlation. If this is not the case, then the cross correlation does not have a peak at the travel time (middle figure). Right figure: The main contribution to the singular components of the cross correlation is from pairs of ray segments issuing from a source \( y \) going to \( x_1 \) and to \( x_2 \), respectively (solid and dashed lines, respectively), and starting in the same direction.

6. Passive sensor imaging with cross correlations. In this section we show using again the stationary phase method that it is possible to image reflectors by cross correlations of signal amplitudes generated by ambient noise sources and recorded by passive sensors. We assume here that the background medium is homogeneous or smoothly varying. Background media with random inhomogeneities and multiple scattering are considered in Section 7.
6.1. Identification of singular components of cross correlations for different configurations of sources sensors and reflectors. The stationary phase analysis in the following subsection indicates that we should distinguish three types of configurations of sources, sensors and reflectors (Figure 6.1). We use terminology from analogous situations in photography but it should be kept in mind that imaging is coherent here, which means that the recorded signals are time-resolved amplitudes and not just intensities.

1) The noise sources surround both the sensors and the reflectors. We call this the surround light configuration.

2) The noise sources are spatially localized and the sensors are between the sources and the reflectors. More precisely, these are configurations in which the rays going through reflectors and sensors reach into the source region, and the sources are between the reflectors and the sensors along these rays. We call this the daylight configuration.

3) The noise sources are spatially localized and the reflectors are between the sources and the sensors. We call this the backlight configuration, in analogy with photography.

We will show that imaging of reflectors can be done by migration of cross correlations. However, the migration imaging functional depends on the configuration of sources, sensors and reflectors. The stationary phase analysis in the next subsection makes clear which is the appropriate imaging functional in each configuration. This can also be done more informally using our understanding from Section 5 of how singular components appear in cross correlations. As noted in that section, the main contributions to the cross correlations $C^{(1)}(\tau, x_1, x_2)$ come from pairs of ray segments starting from a point in the noise source region and going through $x_2$ and $x_1$, respectively, with the same initial angle. When there is a reflector at $z_r$ then we must also consider rays that are reflected at $z_r$. Since the reflectors are assumed to be weak, the important pairs of ray segments are those with one direct ray and one broken ray going from the source region to $x_1$ and $x_2$, respectively. In the daylight configuration, the singular component of $C^{(1)}(\tau, x_1, x_2)$ due to the reflector comes from two pairs of ray segments, as shown in Figure 6.2. From this figure we see that the difference in travel times of the two ray segments is $\pm [\tau(x_2, z_r) + \tau(x_1, z_r)]$, with $+$ for the pair of ray segments in the left figure and $-$ for the pair of ray segments in the right figure. We show in Figure 6.3 two pairs of ray segments that contribute to the singular components of the cross correlation due to a reflector in the backlight configuration. From this figure we see that the differences in travel time of the two ray segments are $\tau(x_2, z_r) - \tau(x_1, z_r)$. 
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Fig. 6.2. Pairs of ray segments that contribute to the singular component of the cross correlation of signals between \(x_1\) and \(x_2\) due to the reflector at \(z_r\), in a daylight configuration. The ray segments start from a point \(y\) in the noise source region with the same angle. One goes to \(x_1\) (solid) and the other one goes to \(x_2\) (dashed).

Fig. 6.3. Same as in Figure 6.2 but in a backlight configuration.

From these considerations we conclude that there is an important difference between daylight and backlight configurations. In daylight configurations the singular components of the cross correlations are concentrated at the total travel time \(\tau(x_2, z_r) + \tau(x_1, z_r)\), while in backlight configurations they are concentrated at the difference travel time \(\tau(x_2, z_r) - \tau(x_1, z_r)\). This determines the form of the migration imaging functional and its resolution.

6.2. Stationary phase analysis of the cross correlation with reflectors.

We carry out the analysis when the background medium is smoothly varying and there is a point reflector at \(z_r\). Since we assume that the reflectors are weak, we use the Born approximation for the Green’s function evaluated at high frequencies (see Appendix C):

\[
\hat{G}(\omega) = \hat{G}_0(\omega) + \omega^2 \sigma_r \hat{G}_0(\omega) \hat{G}_0(\omega) - \tau(x_1, z_r) - \tau(x_2, z_r).
\]  

Here \(\hat{G}_0\) is the Green’s function (5.1) of the background medium, that is, in the absence of reflector, and \(\sigma_r\) is the reflectivity of the reflector.

**Proposition 6.1.** When there is a point reflector at \(z_r\) with small reflectivity \(\sigma_r\) then the cross correlation \(C^{(1)}(\tau, x_1, x_2)\) has two types of singular components, of order \(O(1)\) with respect to \(\sigma_r\) and \(O(\sigma_r)\) as follows.

1) Terms of order \(O(1)\):

If the ray going through \(x_1\) and \(x_2\) extends into the source region, that is, into the support of the function \(\theta\), then there are either one or two singular components at \(\tau = \pm \tau(x_1, x_2)\).

2) Terms of order \(O(\sigma_r)\):

If the ray going through \(x_1\) and \(z_r\) extends into the source region and if \(x_1\) is between \(z_r\) and the sources, then there is a singular component at \(\tau = \tau(x_1, z_r) + \tau(x_2, z_r)\).
If the ray going through \( x_1 \) and \( z_r \) extends into the source region and if \( z_r \) is between \( x_1 \) and the sources, then there is a singular component at \( \tau = \tau(x_2, z_r) - \tau(x_1, z_r) \).

If the ray going through \( x_2 \) and \( z_r \) extends into the source region and if \( x_2 \) is between \( z_r \) and the sources, then there is a singular component at \( \tau = -\tau(x_1, z_r) - \tau(x_2, z_r) \).

If the ray going through \( x_2 \) and \( z_r \) extends into the source region and if \( z_r \) is between \( x_2 \) and the sources, then there is a singular component at \( \tau = \tau(x_2, z_r) - \tau(x_1, z_r) \).

The non-singular components of order \( O(1) \) of the cross correlation function \( C^{(1)}(\tau, x_1, x_2) \) are supported in the interval \((-\tau(x_1, x_2), \tau(x_1, x_2))\).

The components of order \( O(1) \) are the contributions of the direct waves that have not been reflected by the reflector at \( z_r \). To summarize, to leading order in \( \sigma_r \), we have the following:

1) In a daylight imaging configuration, the leading-order singular components are of order \( O(\sigma_r) \) and at \( \tau = \pm(\tau(x_1, z_r) + \tau(x_2, z_r)) \).
2) In a backlight imaging configuration, the leading-order singular component is of order \( O(\sigma_r) \) and at \( \tau = \tau(x_2, z_r) - \tau(x_1, z_r) \).
3) In a surround light imaging configuration, the leading-order singular components are of order \( O(1) \) and at \( \tau = \pm\tau(x_1, x_2) \). They correspond to direct waves that have not been reflected at \( z_r \). There are also singular components of order \( O(\sigma_r) \) at \( \tau = \tau(x_2, z_r) - \tau(x_1, z_r) \) and at \( \tau = \pm(\tau(x_1, z_r) + \tau(x_2, z_r)) \).

**Proof.** Using (6.1) and the WKB approximation, the Green’s function has the form

\[
\hat{G}(\frac{\omega}{\varepsilon}, x_1, x_2) \sim a(x_1, x_2)e^{i\xi y(x_1, x_2)} + \omega^2 a_r(x_1, x_2)e^{i\xi y(x_1, x_2)}. \tag{6.2}
\]

Here \( y(x_1, x_2) \) is the travel time from \( x_1 \) to \( x_2 \) with a reflection at \( z_r \)

\[
\tau(x_1, x_2) = \tau(x_1, z_r) + \tau(z_r, x_2),
\]

and the corresponding amplitudes are \( a(x_1, x_2) \) and \( a_r(x_1, x_2) \). The reflected amplitude \( a_r \) is much smaller than \( a \) since it is proportional to the reflectivity \( \sigma_r \) of the reflector at \( z_r \).

Using (4.12) and the WKB approximation (6.2) of the Green’s function we have, up to terms of order \( \sigma_r \),

\[
C^{(1)}(\tau, x_1, x_2) \simeq C^{(1)}_{dd}(\tau, x_1, x_2) + C^{(1)}_{dr}(\tau, x_1, x_2) + C^{(1)}_{rd}(\tau, x_1, x_2),
\]

with

\[
C^{(1)}_{dd}(\tau, x_1, x_2) = \frac{1}{2\pi} \int dy(\theta(y) \int d\omega \hat{F}(\omega)\pi(x_1, y)a(x_2, y)e^{i\xi y(x_1, x_2)}),
\]

\[
C^{(1)}_{dr}(\tau, x_1, x_2) = \frac{1}{2\pi} \int dy(\theta(y) \int d\omega \omega^2 \hat{F}(\omega)\pi(x_1, y)a_r(x_2, y)e^{i\xi y(x_1, x_2)}),
\]

\[
C^{(1)}_{rd}(\tau, x_1, x_2) = \frac{1}{2\pi} \int dy(\theta(y) \int d\omega \omega^2 \hat{F}(\omega)\pi_r(x_1, y)a(x_2, y)e^{i\xi y(x_1, x_2)}).
\]

Here the subscript \( d \) denotes direct waves and \( r \) reflected waves. The rapid phases are given by

\[
\omega T_{dd}(y) = \omega(\tau(y, x_2) - \tau(y, x_1) - \tau),
\]

\[
\omega T_{dr}(y) = \omega(\tau(y, x_2) - \tau(y, x_1) - \tau) = \omega(\tau(y, x_2) - \tau(y, z_r) - \tau(z_r, x_1) - \tau),
\]

\[
\omega T_{rd}(y) = \omega(\tau_r(y, x_2) - \tau(y, x_1) - \tau) = \omega(\tau(y, z_r) + \tau(z_r, x_2) - \tau(y, x_1) - \tau).
\]
The term \( C^{(1)}_{dd} \) is of the same form as the function \( C^{(1)} \) in the proof of Proposition 5.1. It has singular components only if \( x_1, x_2 \) and \( y \) are on the same ray. These singular components are supported on \( \pm \tau(x_1, x_2) \) and they are of order \( O(1) \).

The dominant contribution to the term \( C^{(1)}_{dr} \) comes from the stationary points \((\omega, y)\) satisfying

\[
\partial_\omega \left( \omega T_{dr}(y) \right) = 0, \quad \nabla_y \left( \omega T_{dr}(y) \right) = 0,
\]

which gives the conditions

\[
\tau(y, x_2) - \tau(y, z_r) - \tau(z_r, x_1) = \tau, \quad \nabla_y \tau(y, x_2) = \nabla_y \tau(y, z_r).
\]

By Lemma B.1 in the appendix, the second condition implies that \( x_2 \) and \( z_r \) are on the same side of a ray issuing from \( y \). If the points are aligned along the ray as \( y \to x_2 \to z_r \), then the first condition is equivalent to \( \tau = -\tau(z_r, x_2) - \tau(z_r, x_1) \). If the points are aligned along the ray as \( y \to z_r \to x_2 \), then the first condition is equivalent to \( \tau = \tau(z_r, x_2) - \tau(z_r, x_1) \). The analysis of the term \( C^{(1)}_{dr} \) goes along the same lines. The terms \( C^{(1)}_{dr} \) and \( C^{(1)}_{rd} \) are both of order \( O(\sigma_r) \).

The last statement of the proposition concerns the support of the non-singular component of the cross-correlation \( C_{dd} \). Even without applying the stationary phase method, information about the support of \( C_{dd} \) can be obtained from the form (6.3) of \( T_{dd}(y) \) and the triangle inequality \(|\tau(y, x_1) - \tau(y, x_2)| \leq \tau(x_1, x_2)\). Therefore, the support of \( C_{dd} \) as a function of \( \tau \) is in the interval \((-\tau(x_1, x_2), \tau(x_1, x_2))\).

### 6.3. Migration imaging of cross correlations

In order to image the reflectors, we first assume that we know the medium, in the sense that the travel times between the sensors and points in the region around the reflectors to be imaged are known. If in particular the medium is homogeneous then \( \tau(x, y) = |x - y|/c_0 \).

The primary data for imaging is the set \( \{C(\tau, x_j, x_j)\} \) of cross correlation of signals recorded at the sensors located at \( x_j, j = 1, \ldots, N \). Even when the signal to noise ratio (SNR) at the sensors is large, this primary data set cannot be used directly for imaging because peaks in the cross correlations due to the reflectors are very weak compared both to the peaks of the direct waves and to the non-singular components due to the directional energy flux. It is therefore necessary to process the cross correlations before migration imaging.

There are two ways to process the primary cross correlations:

1. If data sets \( \{C\} \) and \( \{C_0\} \) with and without the reflectors, respectively, are available then we can compute the differential cross correlations \( \{C - C_0\} \) at the sensors and migrate them. This takes out noise source and sensor location effects regardless of the type of illumination.

2. If only the data set \( \{C\} \) with the reflectors is available then we can mask the part of the cross correlations in the interval between travel times among the sensors. This includes elimination of the main peaks at the travel times between sensors. We refer to the resulting masked cross correlations as coda cross correlations, \( \{C_{\text{coda}}\} \).

The form of the migration imaging functional for both differential and coda cross correlations depends in an essential way on the type of illumination as explained in the next subsection.

We note also that we need to distinguish the two parts of the cross correlations supported by the positive time axis and by the negative time axis:

\[
C^+(\tau, x_j, x_l) = C(\tau, x_j, x_l)1_{(0, \infty)}(\tau), \quad C^-(\tau, x_j, x_l) = C(\tau, x_j, x_l)1_{(-\infty, 0)}(\tau).
\]
This is because $C^+$ is an estimate of the causal Green’s function, while $C^-$ is an estimate of the anti-causal Green’s function. Therefore migration or back propagation imaging is different for $C^-$ and $C^+$. 

6.4. Migration imaging with daylight illumination. We consider first migration of differential cross correlations with daylight illumination. The Kirchhoff migration (KM) imaging functional at a search point $z^S$ is

$$T_{KM}(z^S) = \frac{1}{2\pi} \int d\omega \sum_{j,l=1}^{N} e^{-i\omega[\tau(z^S,x_j) + \tau(z^S,x_l)]} \left[ \hat{C}^+(\omega,x_j,x_l) - \hat{C}^+_0(\omega,x_j,x_l) \right]$$

$$+ \frac{1}{2\pi} \int d\omega \sum_{j,l=1}^{N} e^{i\omega[\tau(z^S,x_j) + \tau(z^S,x_l)]} \left[ \hat{C}^-(-\omega,x_j,x_l) - \hat{C}^-_0(\omega,x_j,x_l) \right]. \quad (6.4)$$

This functional uses the positive and negative parts of the cross correlations, which as already noted correspond to the causal and anti-causal Green’s functions. That is why the "back" propagation in the functional is backward for the positive part, hence the minus sign for the travel times in the exponent, and forward for the negative part, hence the plus sign. It is a consequence of the stationary phase analysis of the previous subsection that the matched filter should be chosen as an exponential at the sum of the travel times $\tau(z^S,x_j) + \tau(z^S,x_l)$.

By a simple change of variables in the second integral the KM imaging functional (6.4) can be written as

$$T_{KM}(z^S) = \frac{1}{2\pi} \int d\omega \sum_{j,l=1}^{N} e^{-i\omega[\tau(z^S,x_j) + \tau(z^S,x_l)]} \left[ \hat{C}^{\text{sym}}(\omega,x_j,x_l) - \hat{C}^{\text{sym}}_0(\omega,x_j,x_l) \right],$$

where $C^{\text{sym}}(\tau,x_j,x_l) = |C(\tau,x_j,x_l) + C(-\tau,x_j,x_l)|1_{(0,\infty)}(\tau)$, and with a similar expression for $C^{\text{sym}}_0(\tau,x_j,x_l)$.

Migration imaging of coda cross correlations with daylight illumination is entirely analogous. The KM imaging functional at a search point $z^S$ is

$$T_{KM}^{\text{coda}}(z^S) = \frac{1}{2\pi} \int d\omega \sum_{j,l=1}^{N} e^{-i\omega[\tau(z^S,x_j) + \tau(z^S,x_l)]} \hat{C}^{\text{coda}}(\omega,x_j,x_l), \quad (6.5)$$

where now $C^{\text{coda}}(\tau,x_j,x_l) = [C(\tau,x_j,x_l) + C(-\tau,x_j,x_l)]1_{(\tau(x_j,x_l),\infty)}(\tau)$. As noted in the previous subsection, it is necessary to remove the central part of the cross correlation function because the contribution of the cross correlation $C_0(\tau,x_j,x_l)$ in the absence of the reflectors is large but concentrated in the interval $(-\tau(x_j,x_l),\tau(x_j,x_l))$.

By removing this part the imaging functional has the same behavior as the one using differential cross correlations. The coda truncation should be done with a small margin beyond the travel time $\tau(x_j,x_l)$, roughly equal to a few times the background speed times the decoherence time. We note that when the sensors form an array then migration with coda cross correlations cannot image reflectors in the shadow area located close to the array, at a distance smaller than its aperture.

The KM imaging functional (6.4) (or (6.5)) has the same form as the KM imaging functional that is used when the reflectors are illuminated by active sensors at $(x_j)_{j=1,...,N}$ and the data set consists of the response matrix $P(t,x_i,x_j)$. The sensor at $x_j$ emits an impulse with a specified bandwidth, the sensor at $x_i$ records the signal.
Fig. 6.4. Numerical simulations for daylight migration imaging with an array of five passive sensors (triangles on the top left figure). The reflector to be imaged (diamond) is located at \((x_r = 0, z_r = 100)\) and is illuminated by noise sources (circles). The coda correlation and the differential correlation of the signals recorded by the two sensors at the ends of the array are shown at the top right figure. The migration imaging functional (6.4) using differential cross correlations is shown at the bottom left figure. Migration imaging with coda cross correlations (6.5) is shown at the bottom right figure. Here the image window in which \(z_S\) varies is \(40 \times 40\) range resolution units around the reflector.

\[
P(t, \mathbf{x}_i, \mathbf{x}_j),\text{ and the KM migration functional is}
\]

\[
I_{KM}(x_S) = \frac{1}{2\pi} \int d\omega \sum_{j,l=1}^N e^{-i\omega[\tau(x_S, \mathbf{x}_j) + \tau(x_S, \mathbf{x}_l)]} \hat{P}(\omega, \mathbf{x}_j, \mathbf{x}_l) .
\]  

(6.6)

It is remarkable that daylight migration imaging with passive sensor cross correlations is essentially equivalent to migration imaging with active sensors.

A resolution analysis of migration imaging with active sensor arrays has been carried out in detail [6, 7]. It applies directly to migration imaging with cross correlations as introduced here. We see that the cross range resolution for a linear sensor array with aperture \(a\) is given by \(\lambda a/L\). Here \(L\) is the distance between the sensor array and the reflectors and \(\lambda\) is the central wavelength. The range resolution is proportional to the background velocity times the inverse of the bandwidth.

In Figure 6.4 we show the results of numerical simulations for a sparse array of 5 sensors and with one reflector. There are 200 point noise sources distributed over the domain \([-50, 50] \times [0, 15]\). The power spectral density of the noise sources is a Gaussian \(\exp(-\omega^2/2)\). The speed and the bandwidth are normalized to one so that the range resolution is also normalized to one. The image window is \(40 \times 40\) range resolutions. It is clear that passive imaging with coda cross correlations gives very
good results.

6.5. Migration imaging with backlight illumination. We consider first migration imaging with backlight illumination when we have differential cross correlation sensor data \( \{ C - C_0 \} \). The KM imaging functional at a search point \( z^S \) is

\[
\mathcal{I}^{KM}(z^S) = \frac{1}{2\pi} \int d\omega \sum_{j,l=1}^N e^{-i\omega[\tau(z^S, x_j) - \tau(z^S, x_l)]} \left[ \hat{C}(\omega, x_j, x_l) - \hat{C}_0(\omega, x_j, x_l) \right].
\]

The signs of the travel times in the exponent of the matched filter are determined by the analysis of Subsection 6.2. It is shown there that the singular component of \((C - C_0)(\tau, x_j, x_l)\) is at \( \tau = \tau(z_r, x_l) - \tau(z_r, x_j) \). A better version of this imaging functional that distinguishes between positive and negative parts of the differential cross correlations is

\[
\mathcal{I}^{KM}(z^S) = \frac{1}{2\pi} \int d\omega \sum_{j,l=1}^N e^{-i\omega[\tau(z^S, x_j) - \tau(z^S, x_l)]} \times \left\{ [\hat{C}^+(\omega, x_j, x_j) - \hat{C}_0^+(\omega, x_j, x_j)] 1_{(0,\infty)}(\tau(z^S, x_j) - \tau(z^S, x_j)) + [\hat{C}^-(\omega, x_j, x_j) - \hat{C}_0^-(\omega, x_j, x_j)] 1_{(-\infty,0)}(\tau(z^S, x_j) - \tau(z^S, x_j)) \right\}.
\]

We note that (6.7) has the same form as the incoherent interferometric imaging functional that is used when \( z^S \) is a source emitting an impulse that is recorded by passive sensors at \( \{ x_j \}_{j=1,...,N} \) and the data is the vector \( P(t, x_j) \). The incoherent interferometric functional (IINT) has the form

\[
\mathcal{I}^{IINT}(z^S) = \frac{1}{2\pi} \int d\omega \left| \sum_{l=1}^N e^{-i\omega\tau(z^S, x_l)} \hat{P}(\omega, x_l) \right|^2 = \frac{1}{2\pi} \int d\omega \sum_{j,l=1}^N e^{-i\omega[\tau(z^S, x_l) - \tau(z^S, x_j)]} \hat{P}(\omega, x_l) \overline{\hat{P}(\omega, x_j)},
\]

which is identical to a matched field imaging functional. A resolution analysis of it has already been carried out [7] and can be directly used here. We see that backlight cross correlation imaging with passive sensor arrays does not provide any range resolution, as in incoherent interferometric imaging. When the sensors are distributed or when the array is large then range resolution can be obtained by geometric triangulation.

Let us consider backlight imaging when we only have data in the presence of reflectors. The main contribution of the reflector in the cross correlation \( C(\tau, x_j, x_l) \) is at \( \tau = \tau(z_r, x_l) - \tau(z_r, x_j) \). It is buried in the interval \( (-\tau(x_j, x_l), \tau(x_j, x_l)) \) that contains the main contributions from the noise sources. This is because of the triangle inequality \( |\tau(z_r, x_l) - \tau(z_r, x_j)| \leq \tau(x_j, x_l) \). Therefore it is impossible to use masking so as to amplify the effect of the reflectors. The only thing that can be done is to back propagate the full cross correlations. The KM imaging functional at a search point \( z^S \) has the form

\[
\mathcal{I}^{KM}(z^S) = \frac{1}{2\pi} \int d\omega \sum_{j,l=1}^N e^{-i\omega[\tau(z^S, x_l) - \tau(z^S, x_j)]} \hat{C}(\omega, x_j, x_l).
\]

We see in Figure 6.5 that the quality of this image is inferior to the one obtained with the functional (6.7) using differential cross correlations.
Fig. 6.5. Numerical simulations for backlight migration imaging with an array of five passive sensors (triangles on the top left figure). The reflector to be imaged (diamond) is located at \((x_r = 0, z_r = 50)\) and is illuminated by noise sources (circles). The coda correlation and the differential correlation of the signals recorded by the two sensors at the ends of the array are shown at the top right figure. The migration imaging functional (6.7) using differential cross correlations is shown at the bottom left figure. It is not possible to use the coda cross correlation here as was done with daylight imaging in Figure 6.4. Now we back propagate the full cross correlation function with formula (6.8) and we get the bottom right figure. The search window is \(80 \times 80\) range resolution units around the reflector.

6.6. Migration imaging with surround light illumination. With surround light illumination the noise sources are distributed everywhere around the passive sensors and the reflectors. The imaging functional (6.4) with differential cross correlations has the same properties as in the one with daylight illumination. Similarly, the imaging functional (6.7) has the same properties as the one with backlight illumination. Therefore both imaging functionals can be used so as to possibly enhance the signal to noise ratio. However the daylight imaging functional (6.4) has a much better range resolution than (6.7) and so it should be used exclusively if possible. In migration imaging with coda cross correlations it is clearly preferable to use the daylight imaging functional (6.5).

We note that when imaging with cross correlations it may also be possible to estimate the propagation speed of the background medium when it is not known but is smoothly varying. The sensor array must in such cases be distributed in a suitable way in the region of interest. We must first use the cross correlations to estimate travel times between sensors. Then we can estimate the background velocity with a least squares tomographic method [3]. Migration imaging can now be done with travel times based on the estimated background velocity.
7. Passive sensor imaging with cross correlations in a scattering medium. In this section we consider imaging by cross correlations, as in the previous section, but in addition to the smoothly varying background speed there are random inhomogeneities that cause multiple scattering. There are both advantages and disadvantages when imaging reflectors in such an environment. One important advantage is that even localized noise sources can appear to have considerably enhanced directional diversity because of the scattering medium. An inevitable disadvantage is loss of resolution and blurring because of the environmental inhomogeneities.

A qualitative way to assess this trade-off is by using the transport mean free path as a measure of distances over which coherent wave propagation transforms into incoherent wave energy transport. Coherent sensor imaging as introduced in the previous section cannot be expected to work unless distances between sensors and reflectors are of the order of the transport mean free path or shorter. It is also expected that the directional diversity of the noise sources will be enhanced if the distances between sensors and reflectors to noise sources are large compared to the transport mean free path. Under these circumstances, coherent sensor imaging will benefit from enhanced directional diversity of the noise sources without too much loss of resolution and blurring.

7.1. Structure of the cross correlation in a scattering medium. There are two different contributions in the empirical cross correlations $C_T(\tau, x_1, x_2)$, in addition to the singular component at $\tau = \pm \tau(x_1, x_2)$ corresponding to the direct waves from the noise sources.

One contribution comes from the residual fluctuations of the noise sources, which emit stationary random signals. The empirical correlations $C_T$ are random functions with standard deviation of order $T^{-1/2}$, as described in Appendix A. These fluctuations decay as the integration time $T$ becomes larger. The other contribution comes from the multiple scattering of waves by the inhomogeneities in the environment. These contributions are there because even though the identity (2.2) does not hold, the cross correlations contain information beyond an estimate of the travel time between the sensors. We expect that for $T$ large, the tails of $C_T$ capture approximately the tails of the Green's function that are generated by the scattered waves. The tails of $C_T$ contain information about both coherent waves, due to the reflectors that we want to image, as well as incoherent superpositions of waves multiply scattered by the random inhomogeneities.

In the remainder of this section we assume that the integration time $T$ is large enough so that the residual fluctuations from the noise sources are negligible compared to those from multiple scattering.

7.2. Iterated cross correlations and coherent interferometry. As in the case of imaging in a smoothly varying medium discussed in Section 6, the sensor cross correlations $\{C(\tau, x_j, x_l), \ j, l = 1, \ldots, N\}$ need to be processed in order to amplify the contribution of the reflectors. We assume here that we have sensor data in the same environment in the absence of the reflectors, and for simplicity we let $C$ stand for differential cross correlations. It is unlikely that imaging in a scattering environment can be done without having differential cross correlations. We need to take out the effects of the environment as much as possible if migration in a fictitious, smoothly varying background is to be at all successful.

We noted in Subsection 6.4, below equation (6.6), that with favorable illumination the passive sensor cross correlations, or differential cross correlations, provide an estimate for the bandlimited impulse response, or differential impulse response, of an
equivalent active sensor network. This is true in a smoothly varying background and in the presence of weak scatterers when the single scattering approximation can be used, as we show in a similar context in Section 8. It is expected to be true even with strong multiple scattering but coherent sensor imaging in such an environment is unlikely to be successful.

Assuming that the passive sensor differential cross correlations provide a good estimate of the equivalent active sensor differential impulse response, we can implement the coherent interferometric imaging techniques (CINT) developed for active sensors [8]. The key idea in CINT is to migrate cross correlations of the impulse response, rather than the impulse response itself. In the present context this translates into migrating cross correlations of the differential cross correlations $C(\tau, x_2, x_1)$, which means that we use fourth-order cross correlations. It is important, however, to compute these fourth order cross correlations locally in time and space, and not over the whole time interval and the whole set of pairs of sensors, as we now explain.

In the time domain, the components of the cross correlations $C(\tau, x_j, x_i)$ associated with the reflectors do not appear as clean peaks at travel times corresponding to those between sensors and reflectors. They have, in addition, codas generated by the inhomogeneities of the scattering medium. The typical duration of such a coda is called delay spread. We segment the cross correlations $C(\tau, x_j, x_i)$ into time intervals and compute their cross correlations locally in each interval. The width $T_d$ of the time intervals should be chosen to be comparable to the delay spread. If $T_d$ is small compared to the delay spread, then the fourth order local cross correlations cannot compress the coda and the images have speckles. If $T_d$ is large compared to the delay spread, then the images are smooth but there is too much loss of resolution and blurring. However, the delay spread is usually not known since it depends on the random medium. It is shown in [7] that the reciprocal of the delay spread is proportional to the decoherence frequency of the Fourier transform of $C(\tau, x_j, x_i)$. Therefore, the delay spread could be estimated by exploiting this remark, and then used to determine the time windows over which local fourth order cross correlations are computed. It is more efficient, and the images have better resolution, if we choose the width $T_d$ of
the time intervals adaptively, as we discuss at the end of this subsection. Given $T_d$, the CINT daylight imaging functional has the form

$$I_{\text{CINT}}(z^S) = \sum_{j,j',l,l'=1}^{N} \int \int_{|\omega - \omega'| \leq T_d^{-1}} \hat{C}(\omega, x_j, x_l) \overline{C}(\omega', x_{j'}, x_{l'}) \times e^{-i\omega|\tau(x_j, z^S) + \tau(x_l, z^S)|} e^{i\omega'|\tau(x_{j'}, z^S) + \tau(x_{l'}, z^S)|} d\omega d\omega' .$$

(7.1)

We can compare this expression to the square of the KM imaging functional (6.4), which has the form

$$|I_{\text{KM}}(z^S)|^2 = \sum_{j,j',l,l'=1}^{N} \int \int \hat{C}(\omega, x_j, x_l) \overline{C}(\omega', x_{j'}, x_{l'}) \times e^{-i\omega|\tau(x_j, z^S) + \tau(x_l, z^S)|} e^{i\omega'|\tau(x_{j'}, z^S) + \tau(x_{l'}, z^S)|} d\omega d\omega' ,$$

(7.2)

where we recall that here $C$ stands for differential cross correlation. We see that the CINT imaging functional (7.1) and the square of the KM functional (7.2) differ only in that the frequencies $|\omega - \omega'| > T_d^{-1}$, where fourth order correlations are weak, are eliminated in CINT.

It is often possible to go one step further in specializing the CINT imaging functional in the case of passive sensor arrays. There is a characteristic decoherence length in the spatial domain, which is similar to the decoherence frequency. This decoherence length is the distance between sensors beyond which the coda that can be recorded at them are not correlated. The cross correlations of such coda tend to increase the fluctuations of the imaging functional. We therefore remove them. The resulting CINT imaging functional at a search point $z^S$ has the form

$$I_{\text{CINT}}(z^S) = \sum_{l,j,l',j'=1}^{N} \int \int_{|\omega - \omega'| \leq T_d^{-1}} d\omega d\omega' \hat{C}(\omega, x_j, x_l) \overline{C}(\omega', x_{j'}, x_{l'}) \times e^{-i\omega|\tau(x_j, z^S) + \tau(x_l, z^S)|} e^{i\omega'|\tau(x_{j'}, z^S) + \tau(x_{l'}, z^S)|} .$$

After a change of variables it can be seen more clearly that the fourth order cross correlations are computed only over nearby frequencies and sensors:

$$I_{\text{CINT}}(z^S) = \sum_{j,j'=1}^{N} \sum_{l,l'=1}^{N} \int \int_{|\omega - \omega'| \leq T_d^{-1}} d\omega d\omega' \hat{C}(\omega + \frac{\Omega}{2}, x_j, x_l) \overline{C}(\omega - \frac{\Omega}{2}, x_{j'}, x_{l'}) \times e^{-i\frac{\Omega}{2}|\tau(x_j, z^S) - \tau(x_l, z^S)|} e^{i\frac{\Omega}{2}|\tau(x_{j'}, z^S) - \tau(x_{l'}, z^S)|} .$$

It is shown in [8] that the range resolution of CINT is of the order of $T_d$ times the background speed and the cross range resolution of the order of $\Delta L / X_d$.

An adaptive procedure for estimating optimally the unknown parameters $T_d$ and $X_d$ is based on minimizing a suitable norm of the image to improve its quality, both in terms of resolution and signal-to-noise ratio. This norm should penalize speckle fluctuations and also should minimize blurring. A possible choice [8] is a sum of the $L^1$ norm of the functional and of the $L^1$ norm of the gradient of the normalized functional $\sqrt{I_{\text{CINT}}(z^S)} / \sup_z \sqrt{I_{\text{CINT}}(z)}$. This normalization is chosen so that when there is no smoothing, $X_d \to \infty$ and $T_d^{-1} \to \infty$, then we recover the KM functional.
8. Iterated cross correlations for travel time estimation in a weakly scattering medium. For travel time tomography to be successful it is necessary to have good estimates of the travel times between pairs of sensors that cover well the region of interest. When the noise sources are spatially localized and there is a strong directional energy flux at the sensors, then travel time estimates will be poor for sensor pairs in directions perpendicular to this flux. In this section we show that it is possible to exploit scattering from random inhomogeneities so as to enhance travel time estimation.

8.1. The use of auxiliary sensors. We consider distributions of noise sources that are spatially localized and media with scattering that is not strong enough for equipartition of the fields at the sensors \cite{34}. Therefore, even with scattering, the signals depend strongly on the spatial localization of the noise sources, which affects the quality of travel time estimation. However, the coda of the cross correlations are generated by scattered waves, which have more directional diversity than the direct waves from the noise sources. By cross correlating the coda of the cross correlations, which produces fourth order cross correlations, it is possible to exploit scattered waves and their enhanced directional diversity. Campillo and Stehly \cite{13} suggest a way to estimate the Green’s function between $x_1$ and $x_2$ as follows.

1) Calculate the cross correlations between $x_1$ and $x_{a,j}$ and between $x_2$ and $x_{a,j}$ for a large number $N$ of auxiliary sensors $(x_{a,j})_{j=1,...,N}$ that are distributed over the medium $C_T(\tau, x_{a,j}, x_l) = \int_{0}^{T} u(t, x_{a,j})u(t + \tau, x_l)dt, \; l = 1, 2$.

2) Cross correlate the coda (i.e. the tails) of these cross correlations and sum them over all auxiliary sensors to form the coda cross correlation between $x_1$ and $x_2$:

$$C_{T'}(\tau, x_1, x_2) = \sum_{j=1}^{N} \int_{-T', -T_{\text{coda}}}^{T'} C_T(\tau', x_{a,j}) C_T(\tau', x_{l}, x_{a,j}) d\tau'.$$

Here the time $T_{\text{coda}}$ is chosen so as to eliminate the coherent part of $C_T$. In general, $T_{\text{coda}}$ should be larger than $\max(\tau(x_{a,j}, x_1), \tau(x_{a,j}, x_2))$.

Using the stationary phase method we show in the next subsections that the algorithm proposed by Campillo and Stehly can exploit the enhanced directivity of scattered waves. We show in particular that the coda cross correlation $C_{T', T}$, defined by (8.1), has singular components at the travel time between the sensors even in the unfavorable case in which the ray joining $x_1$ and $x_2$ does not reach into the source region, as discussed in Section 5.

8.2. A model for the scattering medium and the auxiliary sensors. In order to analyze the coda (fourth-order) cross correlation (8.1) we first introduce a model for the inhomogeneous medium. A simple single-scattering model is sufficient for this purpose. We assume that the medium has of a smoothly varying background speed $c_0(x)$, which we want to image, and a large collection of point scatterers at $(z_{s,j})_{j\geq 1}$. Their reflectivities $\sigma_j$ are assumed to be zero-mean independent random variables with small variance $\sigma^2$. In the Born (single-scattering) approximation (see Appendix C) the full Green’s function at scaled high frequencies is given by

$$\hat{G}(\frac{\omega}{\varepsilon}, x, y) = \hat{G}_0(\frac{\omega}{\varepsilon}, x, y) + \hat{G}_1(\frac{\omega}{\varepsilon}, x, y),$$

(8.2)
where $\hat{G}_0$ is the Green’s function (5.1) of the background medium and $\hat{G}_1$ is given by
\[
\hat{G}_1\left(\frac{\omega}{\varepsilon}, x, y\right) = \omega^2 \sum_j \sigma_j \hat{G}_0\left(\frac{\omega}{\varepsilon}, x, z_{s,j}\right) \hat{G}_0\left(\frac{\omega}{\varepsilon}, z_{s,j}, y\right).
\] (8.3)

We also assume that:

1) The auxiliary sensors at $(x_{a,j})_{j=1,\ldots,N}$ are distributed with a continuum density $\chi(x_a)$. This means that we can approximate sums over the auxiliary sensors by integrals with density $\chi$:
\[
\sum_j \psi(x_{a,j}) \simeq \int dx_a \chi(x_a) \psi(x_a),
\] (8.4)
for any real-valued test function $\psi$.

2) The scatterers at $(z_{s,j})_{j\geq 1}$ are distributed with a continuum density $\rho(z_s)$. Since the reflectivities $\sigma_j$ are zero-mean independent random variables with variance $\sigma^2$, by the central limit theorem we can approximate the distribution of vectors by a Gaussian
\[
\left(\sum_j \sigma_j \psi_1(z_{s,j}), \ldots, \sum_j \sigma_j \psi_n(z_{s,j})\right) \simeq \mathcal{N}(0, Q),
\] (8.5)
for any real-valued test functions $\psi_1, \ldots, \psi_n$. Here $\mathcal{N}(0, Q)$ stands for the multivariate normal distribution with zero-mean and covariance matrix $Q$ given by
\[
Q_{jl} = \sigma^2 \int dz_s \rho(z_s) \psi_j(z_s) \psi_l(z_s).
\]

8.3. Stationary phase analysis for coherent fourth-order cross correlations. We distinguish two components in the cross correlation function $C^{(1)}$ defined by (4.10):

1) The coherent component $C^{(1)}_{\text{coh}}$ is associated with the direct waves that have not
been scattered. From the representation (8.2) of the Green’s function, the coherent component \( C^{(1)}_{\text{coh}} \) of the cross correlation function is given by

\[
C^{(1)}_{\text{coh}}(\tau, x_1, x_2) = \int dy \theta(y) \int ds ds' F^c(s') G_0(s, x_1, y) G_0(\tau + s', x_2, y),
\]

where \( G_0 \) is the Green’s function of the smoothly varying background.

2) The remaining coda component \( C^{(1)}_{\text{coda}} \), which is the contribution of the waves scattered by the inhomogeneities:

\[
C^{(1)}_{\text{coda}}(\tau, x_1, x_2) = C^{(1)}(\tau, x_1, x_2) - C^{(1)}_{\text{coh}}(\tau, x_1, x_2).
\]

We consider first the coherent fourth-order cross correlation function

\[
C^{(2)}(\tau, x_1, x_2) = \lim_{T' \to \infty} \int_{-T'}^{T'} d\tau' \sum_j C^{(1)}_{\text{coh}}(\tau', x_{a,j}, x_1) C^{(1)}_{\text{coh}}(\tau' + \tau, x_{a,j}, x_2)
\]

\[
= \lim_{T' \to \infty} \int_{-T'}^{T'} d\tau' \int dx_a \chi(x_a) C^{(1)}_{\text{coh}}(\tau', x_a, x_1) C^{(1)}_{\text{coh}}(\tau' + \tau, x_a, x_2).
\]

The following proposition is analogous to Proposition 5.1.

**Proposition 8.1.** The cross correlation \( C^{(2)}(\tau, x_1, x_2) \) has singular components if and only if the ray going through \( x_1 \) and \( x_2 \) extends into the source region. If this is the case, there are one or two singular components at \( \tau = \pm \tau(x_1, x_2) \).

**Proof.** We consider the coherent fourth-order cross correlation function and write it as

\[
C^{(2)}(\tau, x_1, x_2) = \int dx_a \chi(x_a) \int d\tau' C^{(1)}_{\text{coh}}(\tau', x_a, x) C^{(1)}_{\text{coh}}(\tau' + \tau, x_a, x_2)
\]

\[
= \int dx_a dy_1 dy_2 \chi(x_a) \theta(y_1) \theta(y_2) \int d\tau' ds ds' d\nu F^c(s') F^c(u')
\]

\[
\times G_0(s, x_a, y_1) G_0(\tau' + s + s', x_1, y_1) G_0(\nu, x_a, y_2) G_0(\tau' + \nu + u', x_2, y_2).
\]

In the Fourier domain it has the form

\[
C^{(2)}(\tau, x_1, x_2) = \frac{1}{2\pi} \int dx_a dy_1 dy_2 \chi(x_a) \theta(y_1) \theta(y_2) \int d\omega \hat{F}^c(\omega) e^{-i\omega \tau}
\]

\[
\times \hat{G}_0(\omega, x_a, y_1) \hat{G}_0(\omega, x_1, y_1) \hat{G}_0(\omega, x_a, y_2) \hat{G}_0(\omega, x_2, y_2).
\]

It is remarkable that all four Green’s functions are evaluated at the same frequency, which comes from the fact that only cross correlation operations are involved. We analyze this expression in the high-frequency regime by using the WKB approximation (5.4) of \( \hat{G}_0 \). We find that \( C^{(2)} \) has singular components (at \( \tau = \pm \tau(x_1, x_2) \)) only if we can find stationary phase configurations in which the source points \( y_1, y_2 \), the auxiliary sensor \( x_a \), and the main sensors \( x_1, x_2 \) are along the same ray. In particular, there must be a ray starting inside the source region and going through \( x_1 \) and \( x_2 \), as in Proposition 5.1. \( \square \)

We see, therefore, that there is no gain in using the function \( C^{(2)} \) in place of \( C^{(1)} \) for travel time estimation.
8.4. Stationary phase analysis for coda (fourth-order) cross correlations. As shown in Proposition 8.1, the singular component of the coherent fourth-order cross correlation $C^{(2)}$ has the same properties as the one of the standard cross correlation $C^{(1)}$. We now consider the coda cross correlation, as suggested by Campillo and Stehly, which is given by

$$C^{(3)}(\tau, x_1, x_2) = \lim_{T' \to \infty} \int_{-T'}^{T'} d\tau' \sum_j C^{(1)}_{\text{coda}}(\tau', x_{a,j}, x_1) C^{(1)}_{\text{coda}}(\tau' + \tau, x_{a,j}, x_2)$$

$$= \lim_{T' \to \infty} \int_{-T'}^{T'} d\tau' \int dx_a \chi(x_a) C^{(1)}_{\text{coda}}(\tau', x_a, x_1) C^{(1)}_{\text{coda}}(\tau' + \tau, x_a, x_2) \quad (8.8)$$

**Proposition 8.2.** There are two (and only two) high-frequency (singular) components in $C^{(3)}$, at times $t = \pm \tau(x_1, x_2)$, if the two following conditions hold:

1) There are scatterers located along the ray going through $x_1$ and $x_2$, but not on the segment between $x_1$ and $x_2$. These scatterers are the basic ones for the enhanced travel time estimation.

2) There are auxiliary sensors located along rays going from the source region (the support of the function $\theta$) to the basic scatterers (see Figure 8.2).

The important point here is that it is not required that the ray joining the sensors $x_1$ and $x_2$ should extend into the source region as in Propositions 5.1 and 8.1. Therefore the configurations for which the coda cross correlation $C^{(3)}$ has a singular component at the travel time between $x_1$ and $x_2$ are much more general than the ones for which the standard cross correlation has a singular component. Propositions 8.1 and 8.2 show that there is no additional information about the travel time in the coherent fourth-order cross correlation, but that there is such information in the coda (fourth-order) cross correlation. Therefore, the coda cross correlation is the one that should be used. The full fourth-order cross correlation function, obtained by cross-correlating the full cross correlation functions $C^{(1)} = C^{(1)}_{\text{coh}} + C^{(1)}_{\text{coda}}$, is the sum of the coherent fourth-order cross correlation $C^{(2)}$, the coda fourth-order cross correlation $C^{(3)}$, and cross correlations between the coherent $C^{(1)}_{\text{coh}}$ and the coda $C^{(1)}_{\text{coda}}$. Using again the stationary phase method, we see that these cross correlations have singular components only at $\tau = \pm \tau(x_1, x_2)$, and only if the ray joining $x_1$ and $x_2$ extends into the source region. Therefore, while the interesting singular component of the coda fourth-order cross correlation is present in the full fourth-order cross correlation, it is buried in many other uninteresting terms, especially when the scattered waves are much weaker than the direct waves. That is why it is much better, in terms of the signal-to-noise ratio, to use only the coda fourth-order cross correlations.

**Proof.** We consider the coda cross correlation $C^{(3)}$ in the form

$$C^{(3)}(\tau, x_1, x_2) = \frac{1}{2\pi} \int dx_a dy_1 dy_2 \chi(x_a) \theta(y_1) \theta(y_2) \int d\omega \hat{F}^\tau(\omega)^2 e^{-i\omega \tau} \times \left( \hat{G}(\omega, x_a, y_1) \overline{\hat{G}(\omega, x_1, y_1)} - \hat{G}_0(\omega, x_a, y_1) \overline{\hat{G}_0(\omega, x_1, y_1)} \right) \times \left( \hat{G}(\omega, x_a, y_2) \overline{\hat{G}(\omega, x_2, y_2)} - \hat{G}_0(\omega, x_a, y_2) \overline{\hat{G}_0(\omega, x_2, y_2)} \right)$$

$$= \frac{1}{2\pi} \int dx_a dy_1 dy_2 \chi(x_a) \theta(y_1) \theta(y_2) \int d\omega \hat{F}^\tau(\omega)^2 e^{-i\omega \tau} \times \left( \hat{G}_0(\omega, x_a, y_1) \overline{\hat{G}_1(\omega, x_1, y_1)} + \hat{G}_1(\omega, x_a, y_1) \overline{\hat{G}_0(\omega, x_1, y_1)} \right) \times \left( \hat{G}_0(\omega, x_a, y_2) \overline{\hat{G}_1(\omega, x_2, y_2)} + \hat{G}_1(\omega, x_a, y_2) \overline{\hat{G}_0(\omega, x_2, y_2)} \right) + O(\sigma^3), \quad (8.9)$$
where $\hat{G}_1$ is given by (8.3). We analyze the expectation $E[C^{(3)}]$ with respect to the random medium. By keeping only the leading order terms in $\sigma$, which are here of order $\sigma^2$, there are four contributions to $E[C^{(3)}]$:

$$C^{(3)}_1(\tau, x_1, x_2) = \frac{\sigma^2}{2\pi} \int dx_3 dy_1 dy_2 d\xi_3 \chi(x_3) \theta(y_1) \theta(y_2) \rho(\xi_3) \int d\omega \omega^4 \hat{F}(\omega)^2 e^{-i\frac{\omega^2}{4}\tau}$$

$$\times \hat{G}_0\left(\frac{\omega}{\varepsilon}, x_3, y_1\right) \hat{G}_0\left(\frac{\omega}{\varepsilon}, x_1, z_3\right) \hat{G}_0\left(\frac{\omega}{\varepsilon}, z_3, y_1\right)$$

$$\times \hat{G}_0\left(\frac{\omega}{\varepsilon}, x_3, y_2\right) \hat{G}_0\left(\frac{\omega}{\varepsilon}, x_2, z_3\right) \hat{G}_0\left(\frac{\omega}{\varepsilon}, z_3, y_2\right),$$

$$C^{(3)}_2(\tau, x_1, x_2) = \frac{\sigma^2}{2\pi} \int dx_3 dy_1 dy_2 d\xi_3 \chi(x_3) \theta(y_1) \theta(y_2) \rho(\xi_3) \int d\omega \omega^4 \hat{F}(\omega)^2 e^{-i\frac{\omega^2}{4}\tau}$$

$$\times \hat{G}_0\left(\frac{\omega}{\varepsilon}, x_3, z_3\right) \hat{G}_0\left(\frac{\omega}{\varepsilon}, z_3, y_1\right) \hat{G}_0\left(\frac{\omega}{\varepsilon}, x_1, y_1\right)$$

$$\times \hat{G}_0\left(\frac{\omega}{\varepsilon}, x_3, y_2\right) \hat{G}_0\left(\frac{\omega}{\varepsilon}, x_2, z_3\right) \hat{G}_0\left(\frac{\omega}{\varepsilon}, z_3, y_2\right),$$

$$C^{(3)}_3(\tau, x_1, x_2) = \frac{\sigma^2}{2\pi} \int dx_3 dy_1 dy_2 d\xi_3 \chi(x_3) \theta(y_1) \theta(y_2) \rho(\xi_3) \int d\omega \omega^4 \hat{F}(\omega)^2 e^{-i\frac{\omega^2}{4}\tau}$$

$$\times \hat{G}_0\left(\frac{\omega}{\varepsilon}, x_3, y_1\right) \hat{G}_0\left(\frac{\omega}{\varepsilon}, x_1, z_3\right) \hat{G}_0\left(\frac{\omega}{\varepsilon}, z_3, y_1\right)$$

$$\times \hat{G}_0\left(\frac{\omega}{\varepsilon}, x_3, y_2\right) \hat{G}_0\left(\frac{\omega}{\varepsilon}, x_2, z_3\right) \hat{G}_0\left(\frac{\omega}{\varepsilon}, z_3, y_2\right),$$

$$C^{(3)}_4(\tau, x_1, x_2) = \frac{\sigma^2}{2\pi} \int dx_3 dy_1 dy_2 d\xi_3 \chi(x_3) \theta(y_1) \theta(y_2) \rho(\xi_3) \int d\omega \omega^4 \hat{F}(\omega)^2 e^{-i\frac{\omega^2}{4}\tau}$$

$$\times \hat{G}_0\left(\frac{\omega}{\varepsilon}, x_3, z_3\right) \hat{G}_0\left(\frac{\omega}{\varepsilon}, z_3, y_1\right) \hat{G}_0\left(\frac{\omega}{\varepsilon}, x_1, y_1\right)$$

$$\times \hat{G}_0\left(\frac{\omega}{\varepsilon}, x_3, z_3\right) \hat{G}_0\left(\frac{\omega}{\varepsilon}, z_3, y_2\right) \hat{G}_0\left(\frac{\omega}{\varepsilon}, x_2, y_2\right).$$

We will analyze each of these four contributions in the high-frequency regime using the stationary phase method.
We consider the first contribution \( C^{(3)}_1 \). Using the WKB approximation (5.4) of the background Green’s function, we obtain:

\[
C^{(3)}_1(\tau, x_1, x_2) = \frac{\sigma^2}{2\pi} \int dx_3 dy_1 dy_2 d\zeta \chi(x_3) \theta(y_1) \theta(y_2) \rho(\zeta) \int d\omega \omega^4 \hat{F}(\omega)^2 \\
\times a(x_3, y_1) \overline{a}(x, z_2) a(x, \zeta) a(\zeta, y_2) e^{i \frac{\pi}{2} \tau(x_3, y_1, y_2, \zeta)}.
\]

The rapid phase is

\[
\omega T(x_3, y_1, y_2, z_2) = \omega \left[ \tau(x_3, y_1) - \tau(x_1, z_2) - \tau(z_2, y_2) + \tau(x_3, z_2) + \tau(z_2, y_2) - \tau(x_1, y_1) \right].
\]

The dominant contribution comes from the stationary phase points satisfying

\[
(\partial_\omega, \nabla x_1, \nabla y_1, \nabla y_2, \nabla z_2)(\omega T(x_3, y_1, y_2, z_2)) = (0, 0, 0, 0),
\]

which give the five conditions

\[
\begin{align*}
\tau(x_3, y_1) &- \tau(x_1, z_2) - \tau(z_2, y_1) - \tau(x_3, y_2) + \tau(x_2, z_2) + \tau(z_2, y_2) = \tau, \\
\nabla x_3 \cdot \tau(x_3, y_1) &= \nabla x_3 \cdot \tau(x_3, y_2), \\
\nabla y_1 \cdot \tau(y_1, x_3) &= \nabla y_1 \cdot \tau(y_1, z_2), \\
\nabla y_2 \cdot \tau(y_2, x_3) &= \nabla y_2 \cdot \tau(y_2, z_2), \\
\nabla z_2 \cdot \tau(z_2, x_3) + \nabla z_2 \cdot \tau(z_2, y_2) &= \nabla z_2 \cdot \tau(z_2, x_2) + \nabla z_2 \cdot \tau(z_2, y_2).
\end{align*}
\]

It turns out that there do exist such stationary points for \( C^{(3)}_1 \). They are determined by using Lemma B.1. The second condition means that \( y_1 \) and \( y_2 \) lie on the same side of a ray issuing from \( x_3 \). The third one means that \( x_3 \) and \( z_2 \) lie on the same side of a ray issuing from \( y_1 \). Therefore, \( y_1, y_2, z_2 \) and \( x_3 \) are on the same ray, with the pairs \( (y_1, y_2) \) and \( (z_2, x_3) \) on opposite sides. There are four possibilities: \( y_1 \rightarrow y_2 \rightarrow z_2 \rightarrow x_3, y_2 \rightarrow y_1 \rightarrow z_2 \rightarrow x_3, y_1 \rightarrow y_2 \rightarrow x_3 \rightarrow z_2, and y_2 \rightarrow y_1 \rightarrow x_3 \rightarrow z_2.\)

The fourth condition is now automatically satisfied and the last condition can be simplified to \( \nabla x_3 \cdot \tau(z_2, x_1) = \nabla x_3 \cdot \tau(z_2, x_2) \), which means that \( x_1 \) and \( x_2 \) are on the same side of a ray issuing from \( z_2 \). There are two possibilities here: \( z_2 \rightarrow x_1 \rightarrow x_2 \) and \( z_2 \rightarrow x_2 \rightarrow x_1 \). Consequently, there are eight types of configurations of points satisfying the last four conditions (Figure 8.3). They all contribute to the term \( C^{(3)}_1 \) if in addition the first condition is satisfied, which is simply \( \tau = \tau(x_1, x_2) \) if \( x_1 \) is between \( z_2 \) and \( x_2 \), and \( \tau = -\tau(x_1, x_2) \) if \( x_2 \) is between \( z_2 \) and \( x_1 \). Therefore, there are only two singular components in \( C^{(3)}_1 \), which are at times \( \tau = \pm \tau(x_1, x_2) \). By considering the Hessian of \( \omega T \) (with respect to \( \omega, x_3, y_1, y_2, z_2 \)) we find that the multiplicative factors given by the stationary phase theorem for these stationary maps are of order \( \varepsilon^{3(d-1)/2} \) (this will be used below).

We examine next the three other terms \( C^{(3)}_j, j = 2, 3, 4 \). They also have stationary points that give peaks at times \( \tau = \pm \tau(x_1, x_2) \), but only in configurations for which a source point \( (y_1, y_2) \), a scatterer point \( (z_2) \) and the two sensors \( x_1 \) and \( x_2 \) are along the same ray, while the other source point, the scatterer \( z_2 \) and the auxiliary sensor \( x_3 \) are along another ray. In particular, a source and the sensors \( x_1 \) and \( x_2 \) must be on the same ray. If this is the case the stationary phase approximation gives a factor of order \( \varepsilon^{3(d-1)/2} \).
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To complete the proof we need to show that the fluctuations of $C^{(3)}$ due to the random medium are small. The calculations carried out for the expectation $\mathbb{E}[C^{(3)}]$ give then correct results for $C^{(3)}$ itself. We analyze the variance Var($C^{(3)}$) = $\mathbb{E}[(C^{(3)})^2] - \mathbb{E}[C^{(3)}]^2$ and show that it is much smaller than $\mathbb{E}[C^{(3)}]^2$. Using (8.9), the property (8.5), and the standard rule for the calculations of fourth-order moments of Gaussian processes, we find that the leading order term of the variance is of order $\sigma^4$ and consists of 32 contributions (there are 48 contributions for $\mathbb{E}[(C^{(3)})^2]$, and among them 16 are canceled by $\mathbb{E}[C^{(3)}]^2$). One of these contributions is

$$V_{41}^{(3)} = \frac{\sigma^4}{4\pi^2} \int d\mathbf{x}_a d\mathbf{y}_1 d\mathbf{y}_2 d\mathbf{z}_s \lambda(x_a) \theta(y_1) \theta(y_2) \rho(z_s) \int d\omega \omega^4 \hat{F}(\omega)^2 e^{-i \frac{\omega}{\epsilon} \tau}$$

$$\times \int d\mathbf{x}_a d\mathbf{y}_1 d\mathbf{y}_2 d\mathbf{z}_s \lambda(x_a) \theta(y_1) \theta(y_2) \rho(z_s) \int d\omega' \omega'^4 \hat{F}(\omega')^2 e^{i \frac{\omega'}{\epsilon} \tau}$$

$$\times \hat{G}_0 \left( \frac{\omega}{\epsilon}, \mathbf{x}_a, \mathbf{y}_1 \right) \hat{G}_0 \left( \frac{\omega}{\epsilon}, \mathbf{x}_1, \mathbf{z}_s \right) \hat{G}_0 \left( \frac{\omega}{\epsilon}, \mathbf{z}_s, \mathbf{y}_1 \right)$$

$$\times \hat{G}_0 \left( \frac{\omega'}{\epsilon}, \mathbf{x}_a, \mathbf{y}_2 \right) \hat{G}_0 \left( \frac{\omega'}{\epsilon}, \mathbf{x}_2, \mathbf{z}_s \right) \hat{G}_0 \left( \frac{\omega'}{\epsilon}, \mathbf{z}_s, \mathbf{y}_2 \right)$$

$$\times \hat{G}_0 \left( \frac{\omega'}{\epsilon}, \mathbf{x}_a, \mathbf{y}_2 \right) \hat{G}_0 \left( \frac{\omega'}{\epsilon}, \mathbf{x}_2, \mathbf{z}_s \right) \hat{G}_0 \left( \frac{\omega'}{\epsilon}, \mathbf{z}_s, \mathbf{y}_2 \right).$$

Using the WKB approximation of the background Green’s function, we obtain an expression in which the rapid phase is

$$\phi = \omega \left[ \tau(x_a, y_1) - \tau(x_1, z_s) - \tau(x_s, y_1) - \tau(x_a, y_2) + \tau(x_2, z_s') + \tau(z_s', y_2) - \tau \right]$$

$$- \omega' \left[ \tau(x_a, y_1') - \tau(x_1, z_s') - \tau(z_s', y_1') - \tau(x_a', y_2') + \tau(x_2, z_s) + \tau(z_s, y_2') + \tau \right].$$

Stationary points for this exist only if all points are along the same ray. Consequently it is necessary that the ray going through $x_1$ and $x_2$ reach into the source region, and
then the factor given by the stationary phase theorem is of order $\varepsilon^{6(d-1)/2}$. This is much smaller than $\varepsilon^{6(d-1)/2}$, which is the order of magnitude of the main contributions of the square of $E[C^{(3)}]$. The other 31 components of the variance of $C^{(3)}$ behave similarly, and some have even smaller contributions. This shows that the fluctuations of $C^{(3)}$ due to the random medium are small when we consider singular components. □

8.5. Statistical stability of fourth-order cross correlation functions. For completeness, we analyze also the role that the fluctuations $C_T - \langle C_T \rangle$, due to the noise sources, play in fourth-order cross correlations. We show that, as may be expected, they are negligible. More precisely, we consider the fluctuations of the cross correlation $C_T$, and we compute the cross correlations between $(x_1, x_{a,j})$ and $(x_2, x_{a,j})$, for all auxiliary sensors $x_{a,j}$, and we study the cross correlation of the fluctuations

$$\tilde{C}_T^2(\tau, x_1, x_2) = \tilde{C}_T(\tau, x_1, x_2) - \langle \tilde{C}_T(\tau, x_1, x_2) \rangle,$$  \hspace{1cm} (8.10)

we compute the cross correlations between $(x_1, x_{a,j})$ and $(x_2, x_{a,j})$, for all auxiliary sensors $x_{a,j}$, and we study the cross correlation of the fluctuations

$$\tilde{C}_T^{(2)}(\tau, x_1, x_2) = \int_{-T'}^{T'} d\tau' \sum_j \tilde{C}_T(\tau', x_{a,j}, x_1) \tilde{C}_T(\tau' + \tau, x_{a,j}, x_2)$$

$$= \int_{-T'}^{T'} d\tau' \int d\chi(x_a) \tilde{C}_T(\tau', x_a, x_1) \tilde{C}_T(\tau' + \tau, x_a, x_2).$$ \hspace{1cm} (8.11)

**Proposition 8.3.** The cross correlation fluctuations $\tilde{C}_T^{(2)}(\tau, x_1, x_2)$ converge to zero in probability as $T, T' \to \infty$.

More precisely, $TC_T^{(2)}(\tau, x_1, x_2)$ is self-averaging as $T, T' \to \infty$:

$$TC_T^{(2)}(\tau, x_1, x_2) \xrightarrow{T, T' \to \infty} \tilde{C}_T^{(2)}(\tau, x_1, x_2).$$

The normalized cross correlation $\tilde{C}_T^{(2)}(\tau, x_1, x_2)$ has a singular component if and only if the ray going through $x_1$ and $x_2$ reaches into the source region. There are then one or two singular components at $\tau = \pm \tau(x_1, x_2)$.

This proposition shows that the fluctuations in the fourth-order cross correlations due to the noise sources decay to zero as $T^{-1}$. Propositions 5.1 and 8.3 show also that $\tilde{C}_T(\tau, x_1, x_2)$ and $C_T(\tau, x_1, x_2)$ carry the same information in their singular components. In other words, the small fluctuations of the cross correlations due to the noise sources have no new information regarding travel times. As shown by Proposition 8.2, it is the tails (or coda) of the cross correlations generated by the inhomogeneities of the medium that contain buried information about travel times, and this information can be extracted by computing the $C^{(3)}$ cross correlation.

8.6. Higher-order cross correlations. We have considered an inhomogeneous medium with weak, randomly distributed point scatterers. Waves scattered by point scatterers are spherical and have no directional memory of the incident waves. This complete isotropization of the scattered field is desirable in order to have good reconstruction of the Green’s function from coda cross correlations. However, enhanced travel time estimation using iterated cross correlations can occur in other scattering media as well. It may be necessary to use higher-order cross correlations in media in which scattering is not isotropic, as we now explain.

A possible way to use higher-order cross-correlations is to iterate further coda cross correlations as follows:
1) Compute the cross correlation between all pairs of sensors, including between pairs of auxiliary sensors.
2) Cross correlate the coda of the cross correlations in order get fourth-order cross correlation functions:

\[ C_{T,T'}(\tau, x_j, x_{a,l}) = \sum_{k=1}^{N} \int \left[ \int_{-T', T_{\text{coda}}} C_T(\tau', x_{a,k}, x_j) C_T(\tau + \tau', x_{a,k}, x_{a,l}) d\tau' \right] d\tau' \]

for \( j = 1, 2 \) and for all auxiliary sensors at \( x_{a,l}, l = 1, \ldots, N \).
3) Cross correlate the coda of the fourth-order cross correlation functions in order to get the eighth-order cross correlation function:

\[ C_{T,T',T''}(\tau, x_1, x_2) = \sum_{l=1}^{N} \int \left[ \int_{-T'', T_{\text{coda}}} C_{T,T'}(\tau', x_1, x_{a,l}) \right] \times C_{T,T'}(\tau' + \tau, x_2, x_{a,l}) d\tau' \]

The coda cross correlation \( C_{T,T'}(\tau, x_1, x_2) \) proposed by Campillo and Stehly and studied in this paper removes contributions of the direct waves and exploits the scattered waves. Higher-order cross correlations, such as \( C_{T,T',T''}(\tau, x_1, x_2) \) remove contributions of direct waves and of those from single scattering. They exploit mainly multiply-scattered waves. When single scattering is not isotropic then multiply scattered waves tend to be more isotropic than direct and single-scattered waves. When single scattering is essentially isotropic then it is unlikely that higher-order cross correlations will give better results. A more quantitative analysis, including signal-to-noise ratios, is needed in order to substantiate these comments.

9. Summary and conclusions. We have presented and analyzed an integrated framework for reflector imaging and travel time estimation with passive sensor networks, using cross correlations of signals generated by ambient noise sources.

We have shown that imaging of reflectors by migrating cross correlations can be effective provided that it is done appropriately, depending on the positions of the reflectors to be imaged relative to the sensors and the noise sources. More precisely:
1) In daylight configurations (Subsection 6.3), we have shown that the imaging functional is similar to Kirchhoff or travel time migration for passive reflectors illuminated by active sensors. It is, therefore, possible to assess resolution rather well in daylight imaging. In particular, the range resolution is inversely proportional to the bandwidth and the cross range resolution is proportional to the Rayleigh resolution formula for sensor arrays (Subsection 6.4). We can also consider optimal illumination in the cross correlation context following the ideas proposed in [8, 9] in the active array context. This means that weights could be introduced in the daylight imaging functional to compensate for possible imbalances between the illuminations of the different sensors from the ambient noise sources.
2) In backlight configurations (Subsection 6.5), we have shown that the imaging functional is similar to Kirchhoff migration for imaging active sources recorded by passive sensors. Because of this analogy, we know that range resolution is very poor and that only the direction of arrival can be estimated with sensor arrays.
3) We have shown (Section 7) how coherent interferometric imaging (CINT) could be implemented when there is scattering in the medium, and how background velocity estimation, which is necessary for migration, could be carried out using the same passive sensor network.
We have also analyzed travel time estimation by cross correlation of noisy signals. We have shown analytically how directionality of the energy flux from the noise sources affects the quality of the travel time estimates. We have shown how iterated cross correlations with auxiliary sensors in a scattering medium can enhance the quality of the travel time estimates (Section 8). We have identified configurations of scatterer and the auxiliary sensor locations that play a significant role in this enhancement:
1) Scatterers that are along the ray joining the principal sensors.
2) Auxiliary sensors that are along rays connecting source points and scatterers identified in 1).

As noted in the introduction, most applications of cross correlation techniques with passive sensor data are, so far, in geophysics. They may, however, be useful in other contexts, in microwave regimes in particular. Wireless cell phones in an urban environment, or in an indoor environment, could be considered as statistically stationary, spatially distributed noise sources. At a central frequency of 2.4 GHz (12 centimeter wavelength) and a bandwidth of 20 MHz the expected travel time resolution is about 10 meters. With a 200 MHz bandwidth the expected resolution is about 1 meter. Exploiting multiple scattering could make possible the use of passive microwave sensor networks to monitor changes in the environment with differential cross correlations. This would have the advantage of being both passive and rather sensitive but will likely require additional information about the environment, such as indoor floor plans or the location of buildings in an urban setting. A detailed analysis, taking into account the relatively small bandwidth and the scattering properties of microwaves in such environments, is currently under study.

Acknowledgments. The work of G. Papanicolaou was partially supported by US Army grant W911NF-07-2-0027-1, ONR grant N00014-02-1-0088, and grant AFOSR FA9550-08-1-0089.

Appendix A. Proof of Proposition 4.1. In this appendix we prove Proposition 4.1. It establishes statistical stability of the cross correlations with respect to noise source fluctuations.

By the stationarity of the process \( n \), the product \( u(t, x_1)u(t + \tau, x_2) \) is itself a stationary random process in \( t \). Therefore the mean of \( C_T \) is independent of \( T \) and is given by

\[
\langle C_T(\tau, x_1, x_2) \rangle = \langle u(0, x_1)u(\tau, x_2) \rangle.
\]

Using (4.7) we get the following integral representation for the average of the cross correlation function:

\[
\langle C_T(\tau, x_1, x_2) \rangle = \int dy_1 dy_2 \int ds ds' G(s, x_1, y_1)G(s', x_2, y_2) \langle n(-s, y_1) n(\tau - s', y_2) \rangle.
\]

Using the form (4.2) of the autocorrelation function of the sources, we obtain

\[
\langle C_T(\tau, x_1, x_2) \rangle = \int dy_1 dy_2 \int ds ds' G(s, x_1, y_1)G(s + s', x_2, y_2) F_{\epsilon}(s') \Gamma(y_1, y_2).
\]

Using the spatial delta-correlation property (4.6) we obtain (4.11), which in the Fourier domain is (4.12).

In order to prove the self-averaging property of the cross correlation function \( C_T \) we compute its variance. The principle of this computation is the following one. We first write the covariance function of \( C_T \) as a multiple integral which involves
the fourth-order moment of the random process \( n \). Since \( n \) is Gaussian, this fourth-order moment can be written as the sum of products of second-order moments, which makes the computation tractable. Using (4.7) and (4.9), the complete expression of the covariance function is

\[
\text{Cov}(C_T(\tau, x_1, x_2), C_T(\tau + \Delta \tau, x_1, x_2)) = \frac{1}{T^2} \int_0^T \int_0^T dt d\tau' \int ds ds' d\tau d\tau' d\omega d\omega' C(\omega) \tilde{C}(\omega') e^{i\omega' \Delta \tau}
\]

\[
\times G(\omega, x_1, y_1) \hat{G}(\omega', x_2, y_1) \wedge G(\omega, x_1, y_2) \tilde{G}(\omega', x_2, y_2) \text{sinc}^2 \left( \frac{(\omega - \omega') T}{2} \right) e^{i\omega s - i\omega' u}.
\]

The product of second-order moments of the random process \( n \) is

\[
\langle n^2(t - s, y_1) n^2(t - u, y_2) \rangle = F^2(s - u) \delta(y_1 - y_2) \delta(y_1' - y_2').
\]

The fourth-order moment of the Gaussian random process \( n \) is

\[
\langle n^2(t - s, y_1) n^2(t - u, y_2) \rangle \langle n^2(t' - s', y_1') n^2(t' - u', y_2') \rangle = \int_0^T \int_0^T dt d\tau' \int ds ds' d\tau d\tau' d\omega d\omega' C(\omega) \tilde{C}(\omega') e^{i\omega' \Delta \tau}
\]

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Taking the limit $T \to \infty$, we see that the variance is of order $1/T$:

$$T \text{Var}(C_T(\tau, x_1, x_2)) \xrightarrow{T \to \infty} \frac{1}{2\pi} \int d\omega \hat{F}^r(\omega)^2 \left| \int dy_1 \theta(y_1) \tilde{G}(\omega, x_1, y_1) \tilde{G}(\omega, x_2, y_1) \right|^2 + \frac{1}{2\pi} \int d\omega \hat{F}^\varphi(\omega)^2 e^{2i\omega \tau} \left| \int dy_1 \theta(y_1) \tilde{G}(\omega, x_1, y_1) \tilde{G}(\omega, x_2, y_1) \right|^2,$$

which quantifies the convergence rate in (4.13). Note that the first term of the asymptotic variance does not depend on $\tau$, which means that it corresponds to fluctuations for the cross correlation around its mean $C^{(1)}$ that are stationary and extend over the whole time axis. The second term corresponds to local fluctuations. The time scale of the fluctuations of the cross correlation can be quantified from the asymptotic covariance function

$$T \text{Cov}(C_T(\tau, x_1, x_2), C_T(\tau + \Delta \tau, x_1, x_2)) \xrightarrow{T \to \infty} \frac{1}{2\pi} \int d\omega \hat{F}^r(\omega)^2 e^{i\omega \Delta \tau} \left| \int dy_1 \theta(y_1) \tilde{G}(\omega, x_1, y_1) \tilde{G}(\omega, x_2, y_1) \right|^2 + \frac{1}{2\pi} \int d\omega \hat{F}^\varphi(\omega)^2 e^{i\omega \Delta \tau e^{2i\omega \tau}} \left| \int dy_1 \theta(y_1) \tilde{G}(\omega, x_1, y_1) \tilde{G}(\omega, x_2, y_1) \right|^2,$$

(A.4)

which shows that the decoherence time of the fluctuations of $C_T$ is proportional to the decoherence time of the sources.

Appendix B. A remark on the ray equations. The rays are solutions of Hamilton’s equations [6]

$$\frac{dX_t}{dt} = c_0^2(X_t)\xi_t, \quad X_0(x, \xi) = x,$$
$$\frac{d\xi_t}{dt} = -\frac{1}{2} \nabla |c_0|^2(X_t)|\xi_t|^2, \quad \xi_0(x, \xi) = \xi,$$

which derives from the Hamiltonian $H(x, \xi) = \frac{1}{2} c_0^2(x)|\xi|^2$. The quantity $|\xi_0 c_0(X_t) = |\xi| c_0(x) = 1$ is constant along the ray. We assume that any pair of points in the region of interest is connected with only one ray. In other words, for any starting point $y$ and any sensor $x$, there exists a unique vector $\xi$ with norm $|\xi| c_0(y) = 1$ such that $X_t(y, \xi) = x$ at some time $t$, and then the travel time $\tau(x, y)$ is equal to this time $t$.

Lemma B.1. 1. If $\nabla_y \tau(y, x_1) = -\nabla_y \tau(y, x_2)$, then $x_1$ and $x_2$ lie on the same side on the same ray issuing from $y$, and

$$|\tau(y, x_1) - \tau(y, x_2)| = \tau(x_1, x_2).$$

2. If $\nabla_y \tau(y, x_1) = \nabla_y \tau(y, x_2)$, then $x_1$ and $x_2$ lie on the opposite sides of the same ray issuing from $y$, and

$$\tau(y, x_1) + \tau(y, x_2) = \tau(x_1, x_2).$$

Proof. The travel time from $y$ to $x_1$ is associated to a unique ray. We can first look at this ray as starting from $y$, and denote by $\xi_0$ the initial angle vector of this ray, starting from $y$. We can also look at the same ray starting from $x_1$, and we denote by $\xi_1$ the initial angle vector of this ray starting from $x_1$ (see Figure B.1). We have

$$X_t(y, \xi_0) = X_{\tau-t}(x_1, \xi_1), \quad \xi_t(y, \xi_0) = -\xi_{\tau-t}(x_1, \xi_1).$$
Fig. B.1. The ray connecting \( y \) and \( x_1 \) is the ray starting from \( y \) with the angle vector \( \xi_0 \) or equivalently the ray starting from \( x_1 \) with the angle vector \( \xi_1 \).

For any time \( t \), the angle vector is given by

\[
\xi_t(x_1, \xi_1) = \nabla_2 \tau(x_1, X_t(x_1, \xi_1)),
\]

where the gradient is taken with respect to the second point. Applying this identity at time \( t = \tau(x_1, y) \):

\[
-\xi_0 = \xi_\tau(x_1, \xi_1) = \nabla_2 \tau(x_1, X_\tau(x_1, \xi_1)) = \nabla_y \tau(y, x_1).
\]

If the two points \( x_1 \) and \( x_2 \) are such that \( \nabla_y \tau(y, x_1) = \nabla_y \tau(y, x_2) \), then the last equality implies that the ray connecting \( y \) to \( x_1 \) and the ray connecting \( y \) to \( x_2 \) must start with the same angle. This means that \( x_1 \) and \( x_2 \) lie on the same ray issuing from \( y \).

Appendix C. The single-scattering approximation. We describe briefly the single-scattering model used in Subsection 6.2 to calculate cross correlations in the presence of reflectors, and in Section 8 to analyze coda cross correlations. We assume that the medium is characterized by a smoothly varying background velocity \( c_0(x) \) and a set of localized scatterers around \( (z_j)_{j \geq 1} \). We model these scatterers as fluctuations \( V(x) \) in the propagation speed as follows:

\[
\frac{1}{c^2(x)} = \frac{1}{c_0^2(x)} + V(x), \quad V(x) = \sum_j \frac{1}{c_j^2} B_j(x - z_j).
\]

Here \( B_j \) are compactly supported domains. The solution of the full wave equation (4.1) can be written as a sum

\[
u(t, x) = u_0(t, x) + u_1(t, x),\]

where the direct field \( u_0 \) satisfies the wave equation (4.1) with the background velocity \( c_0(x) \) and it is given by:

\[
u_0(t, x) = \int \int G_0(t - s, x, y) n(s, y) ds dy, \quad (C.1)
\]

where \( G_0 \) is the Green’s function of the background medium whose Fourier transform is solution of (5.1). The scattered field \( u_1 \) is given by

\[
u_1(t, x) = \int \int G_0(t - s, x, y) V(y) \frac{\partial^2 u_0}{\partial s^2}(s, y) ds dy.
\]

This expression is exact. The Born approximation (or single scattering approximation) consists in replacing \( u \) on the right side by the field \( u_0 \) [10]:

\[
u_1(t, x) = \int \int G_0(t - s, x, y) V(y) \frac{\partial^2 u_0}{\partial s^2}(s, y) ds dy. \quad (C.2)
\]
This approximation is valid if the scattered field $u_1$ is small compared to the incident field $u_0$. We assume that the diameters of the scattering regions $B_j$ are small compared to typical wavelength and that the velocity contrasts are such that $c_j^{-2} \text{Vol}(B_j) = \varepsilon^2 \sigma_j$, with $\sigma_j$ small. We can then model the scatterers by point scatterers with scattering volumes $\varepsilon^2 \sigma_j$:

$$\frac{1}{c_j^2} 1_{B_j}(x - z_j) \approx \varepsilon^2 \sigma_j \delta(x - z_j).$$

Substituting (C.1) into (C.2) we write the scattered field $u_1$ and the total field $u$ in the form

$$u_1(t, x) = \int G_1(t - s, x, y) n(s, y) ds dy,$$

$$u(t, x) = \int G(t - s, x, y) n(s, y) ds dy,$$

where the full Green’s function $G = G_0 + G_1$ and $G_1$ is given by

$$\tilde{G}_1(\omega, x, y) = \varepsilon^2 \omega^2 \sum_j \sigma_j \tilde{G}_0(\omega, x, z_j) \tilde{G}_0(\omega, z_j, y).$$

**Appendix D. Proof of Proposition 8.3.** We first write the complete expression for the expectation

$$\langle C^{(2)}_{T, T'}(\tau, x_1, x_2) \rangle = \int_{-T'}^{T'} d\tau' \int_{-T}^{T} dt dt' \int ds ds' du du' \int dy_1 dy_1' dy_2 dy_2' dx_a$$

$$\times \chi(x_a) G(s, x_a, y_1) G(u - \tau', x_1, y_2) G(s', x_a, y_1') G(u' - \tau' - \tau, x_2, y_2')$$

$$\times \left( \langle n^x(t - s, y_1) n^x(t - u, y_2) n^x(t' - s', y_1') n^x(t' - u', y_2') \rangle - \langle n^x(t - s, y_1) n^x(t - u, y_2) \rangle \langle n^x(t' - s', y_1') n^x(t' - u', y_2') \rangle \right).$$

(D.1)

We now use the fact that the fourth-order moments of Gaussian processes can be expanded as sums of products of second-order moments, as in the proof of Proposition 4.1. Substituting (A.2) into (D.1) we see that the limit as $T, T' \to \infty$ of $T \langle \tilde{C}^{(2)}_{T, T'} \rangle$ is equal to

$$\tilde{C}^{(2)}(\tau, x_1, x_2) = \frac{1}{2\pi} \int dy_1 dy_2 dx_a \chi(x_a) \theta(y_1) \theta(y_2) \int d\omega \tilde{F}^x(\omega)^2$$

$$\times \tilde{G}(\omega, x_a, y_1) \tilde{G}(\omega, x_1, y_2) \tilde{G}(\omega, x_a, y_1) \tilde{G}(\omega, x_2, y_2) e^{-i\omega \tau}.$$

If we keep only the leading order terms in $\sigma$, which are here of order 1, then we obtain

$$\tilde{C}^{(2)}(\tau, x_1, x_2) = \frac{1}{2\pi} \int dy_1 dy_2 dx_a \chi(x_a) \theta(y_1) \theta(y_2) \int d\omega \tilde{F}(\omega)^2$$

$$\times \tilde{G}_0(\frac{\omega}{\varepsilon}, x_a, y_1) \tilde{G}_0(\frac{\omega}{\varepsilon}, x_1, y_2) \tilde{G}_0(\frac{\omega}{\varepsilon}, x_a, y_1) \tilde{G}_0(\frac{\omega}{\varepsilon}, x_2, y_2) e^{-i\omega \tau} + 0(\sigma).$$

Using the WKB approximation (5.4) of the background Green’s function, we obtain:

$$\tilde{C}^{(2)}(\tau, x_1, x_2) = \frac{1}{2\pi} \int dy_1 dy_2 dx_a \chi(x_a) \theta(y_1) \theta(y_2) \int d\omega \tilde{F}(\omega)^2$$

$$\times \tilde{\pi}(x_a, y_1) a(x_1, y_2) a(x_a, y_1) \tilde{\pi}(x_2, y_2) e^{i\omega \tau}(y_2),$$

$$\tilde{\pi}(x, y) = \int dy_1 dy_2 dx_a \chi(x_a) \theta(y_1) \theta(y_2) \int d\omega \tilde{F}(\omega)^2$$

$$\times \tilde{G}_0(\frac{\omega}{\varepsilon}, x_a, y_1) \tilde{G}_0(\frac{\omega}{\varepsilon}, x_1, y_2) \tilde{G}_0(\frac{\omega}{\varepsilon}, x_a, y_1) \tilde{G}_0(\frac{\omega}{\varepsilon}, x_2, y_2) e^{-i\omega \tau} + 0(\sigma).$$
where the rapid phase is
\[ \omega T(y_2) = \omega [\tau(x_1, y_2) - \tau(x_2, y_2) - \tau]. \]

The analysis of the stationary points of the rapid phase is the same as in the proof of Proposition 5.1, since the rapid phase has the same expression. There are stationary points only if the points \( y_2, x_1 \) and \( x_2 \) are along the same ray.

Thus, the proof of statistical stability goes along the same lines as the one of Proposition 4.1 (Appendix A), using the decomposition of the eighth-order moments of Gaussian processes as sums of products of second-order moments.

REFERENCES